

# SÉMINAIRE DE PROBABILITÉS (STRASBOURG)

HIROSHI TANAKA

## **Limit distribution for 1-dimensional diffusion in a reflected brownian medium**

*Séminaire de probabilités (Strasbourg)*, tome 21 (1987), p. 246-261

[http://www.numdam.org/item?id=SPS\\_1987\\_\\_21\\_\\_246\\_0](http://www.numdam.org/item?id=SPS_1987__21__246_0)

© Springer-Verlag, Berlin Heidelberg New York, 1987, tous droits réservés.

L'accès aux archives du séminaire de probabilités (Strasbourg) (<http://portail.mathdoc.fr/SemProba/>) implique l'accord avec les conditions générales d'utilisation (<http://www.numdam.org/conditions>). Toute utilisation commerciale ou impression systématique est constitutive d'une infraction pénale. Toute copie ou impression de ce fichier doit contenir la présente mention de copyright.

NUMDAM

Article numérisé dans le cadre du programme  
Numérisation de documents anciens mathématiques  
<http://www.numdam.org/>

# LIMIT DISTRIBUTION FOR 1-DIMENSIONAL DIFFUSION IN A REFLECTED BROWNIAN MEDIUM

By H. Tanaka

## Introduction

In analogy with Sinai's problem [8] on a random walk in a random medium, Brox [1] considered the diffusion process  $X(t)$  described by the stochastic differential equation

$$(1) \quad dX(t) = dB(t) - \frac{1}{2} W'(X(t))dt, \quad X(0) = 0,$$

where  $\{W(x), x \in \mathbb{R}\}$  is a Brownian medium independent of another Brownian motion  $B(t)$ , and proved that  $(\log t)^{-2}X(t)$  converges in distribution as  $t \rightarrow \infty$ . Similar results in the case of a considerably wider class of self-similar random media were obtained by Schumacher [7]. Recently Kesten [5] obtained the exact form of the limit distribution for Sinai's random walk as well as for a diffusion in a Brownian medium. See also [2] for a related problem.

In this paper we substitute  $W(x)$  in (1) by a nonnegative reflected Brownian medium and find the corresponding limit distribution. The result was already announced in [9] without proof but the Laplace transform of the limit distribution given in [9: §3] is not correct. We give here a full proof to the whole result of [9: §3] with a correction (see Theorem 1 and 2 below). Our method is similar to that of [1].

Theorem 1. Let  $X(t)$  be a solution of (1) where  $W_+ = \{W(x), x \geq 0\}$  and  $W_- = \{W(-x), x \geq 0\}$  are independent reflected Brownian motions on the half line  $[0, \infty)$  starting from 0 which are also independent of the Brownian motion  $B(t)$ . Then the distribution of  $(\log t)^{-2}X(t)$  converges as  $t \rightarrow \infty$  to the distribution  $\mu$  defined by

$$(2) \quad \mu = \int m_W Q(dW)$$

where  $m_W$  is the probability measure on  $\mathbb{R}$  defined by (3.1) and  $Q$  is the probability measure on the space of media  $W = C(\mathbb{R} \rightarrow 0, \infty) \cap \{W: W(0)=0\}$  such that  $W_{\pm}$  are independent reflected Brownian motions on  $[0, \infty)$ .

Theorem 2.  $\mu$  has a symmetric density and for  $\lambda > 0$

$$(3) \quad \int_0^{\infty} e^{-\lambda x} \mu(dx) = \int_0^{\infty} \frac{\sinh \sqrt{2\lambda}}{\sqrt{2\lambda} \cosh \sqrt{2\lambda} + t \sinh \sqrt{2\lambda}} \cdot \frac{dt}{(1+t)^2}.$$

The present case is not contained in the framework of [7] since the nonnegative reflected medium  $W(x)$  has (uncountably) many points giving its minimum. The case of a nonpositive reflected Brownian medium was discussed in [9]. Some generalizations will be discussed in [5].

Acknowledgment. I wish to thank Professor H. Kesten for pointing out mistakes of the first version of this paper.

### §1. Preliminaries and exit times from valleys

Let  $W$  and  $Q$  be defined as in Theorem 1. For each  $W \in \mathbb{W}$  solutions of the stochastic differential equation (1) define a diffusion process in  $\mathbb{R}$  with generator

$$(1.1) \quad \frac{1}{2} e^{W(x)} \frac{d}{dx} (e^{-W(x)} \frac{d}{dx}) .$$

Such a diffusion can be constructed from a Brownian motion  $B(t)$ <sup>1)</sup> as follows ([4]). Let  $\Omega$  be the space of continuous functions  $\omega : [0, \infty) \rightarrow \mathbb{R}$  with  $\omega(0) = 0$ , and denote by  $P$  the Wiener measure on  $\Omega$ . Denote the value of  $\omega$  at time  $t$  by  $\omega(t)$  or by  $B(t)$  and put

$$L(t, x) = \lim_{\varepsilon \downarrow 0} \frac{1}{\varepsilon} \int_0^t \mathbb{I}_{[x, x+\varepsilon)}(B(s)) ds \quad (\text{local time}),$$

$$S(x) = \int_0^x e^{W(y)} dy ,$$

$$A(t) = \int_0^t e^{-2W(S^{-1}(B(s)))} ds = \int_{\mathbb{R}} e^{-2W(S^{-1}(x))} L(t, x) dx , \quad t \geq 0 ,$$

$S^{-1}, A^{-1}$  = the inverse functions .

Then the process  $X(t, W) = S^{-1}(B(A^{-1}(t)))$  defined on the probability space  $(\Omega, P)$  is a diffusion process with generator (1.1) starting at 0. If we set  $(W^x)(\cdot) = W(\cdot + x)$ , then  $X^x(t, W) = x + X(t, W^x)$  is a diffusion process with generator (1.1) starting at  $x$ . Let

$$T(x_1, x_2) = \inf \{t \geq 0 : B(t) \notin (x_1, x_2)\} ,$$

---

1) The Brownian motion here is not the same as the one in (1) but we use the same notation  $B(t)$  .

$$L(x_1, x_2, x) = L(T(x_1, x_2), x), \quad x \in \mathbb{R},$$

$$S_\lambda(x) = \int_0^x e^{\lambda W(y)} dy,$$

$$X_\lambda(t) = X(t, \lambda W), \quad X_\lambda^x(t) = x + X(t, \lambda W^x).$$

Next we define a valley. Given  $W \in \mathbb{W}$ , a quartet  $V = (a, b_1, b_2, c)$  is called a valley of  $W$  if

- (i)  $a < b_1 < 0 < b_2 < c$ ,
- (ii)  $W(b_1) = W(b_2) = 0$ ,  $W(a) = W(c) = D$ ,
- (iii)  $0 < W(x) < W(a)$  for  $a < x < b_1$ ,  
 $0 < W(x) < W(c)$  for  $b_2 < x < c$ ,
- (iv)  $A_- = \sup \{W(y) - W(x) : a < x < y < b_2\} < D$ ,  
 $A_+ = \sup \{W(x) - W(y) : b_1 < x < y < c\} < D$ .

We call  $D$  (resp.  $A = A_- \vee A_+$ ) the depth (resp. the inner directed ascent) of  $V$ . It is clear that there exist valleys of  $W$  with  $A < 1 < D$  for almost all reflected Brownian media  $W$ .

In what follows let  $W \in \mathbb{W}$  be given and  $V = (a, b_1, b_2, c)$  be a valley of  $W$  with the depth  $D$  and the inner directed ascent  $A$ . We put

$$T_\lambda^x = T_\lambda^x(a, c) = \inf \{t \geq 0 : X_\lambda^x(t) \notin (a, c)\}.$$

The following three lemmas were proved in [1].

Lemma 1. For  $a < x < c$

$$T_\lambda^x(a, c) \stackrel{d}{=} \int_a^c L(\hat{S}_\lambda(a), \hat{S}_\lambda(c), \hat{S}_\lambda(y)) e^{-\lambda W(y)} dy,$$

where

$$\hat{S}_\lambda(y) = \int_x^y e^{\lambda W(z)} dz$$

and  $\stackrel{d}{=}$  means the equality in distribution.

Lemma 2. For each  $\lambda > 0$

$$\{L(\lambda x_1, \lambda x_2, \lambda x), x \in \mathbb{R}\} \stackrel{d}{=} \{\lambda L(x_1, x_2, x), x \in \mathbb{R}\}.$$

---

2)  $a \vee b = \max \{a, b\}$ ,  $a \wedge b = \min \{a, b\}$ .

Lemma 3. For  $\lambda > 0$  and  $W \in \mathbb{W}$

$$(1.2) \quad \{X(t, \lambda W_\lambda), t \geq 0\} \stackrel{d}{=} \{\lambda^{-2} X(\lambda^4 t, W), t \geq 0\},$$

where  $W_\lambda (\in \mathbb{W})$  is defined by

$$W_\lambda(x) = \lambda^{-1} W(\lambda^2 x), \quad x \in \mathbb{R}.$$

The following lemma plays an essential role in our discussions.

Lemma 4. For any  $\lambda > 0$  and  $[u, v] \subset (a, c)$

$$\inf_{u \leq x \leq v} P \left\{ e^{\lambda(D-\delta)} < T_\lambda^x < e^{\lambda(D+\delta)} \right\} \rightarrow 1, \quad \lambda \rightarrow \infty.$$

Proof. The proof is similar to that of the corresponding lemma of [1] but even much simpler. Let  $x \in [u, v]$  be fixed. Setting

$$s_\lambda(y) = \widehat{S}_\lambda(y) / \widehat{S}_\lambda(c) = \int_x^y e^{\lambda W(z)} dz / \int_x^c e^{\lambda W(z)} dz$$

and applying Lemma 1 and 2, we have

$$T_\lambda^x \stackrel{d}{=} \widehat{S}_\lambda(c) \int_a^c L(s_\lambda(a), 1, s_\lambda(y)) e^{-\lambda W(y)} dy.$$

Since

$$\begin{aligned} \widehat{S}_\lambda(c) &\leq (c - x) \exp \left\{ \lambda \max_{[x, c]} W \right\} \quad 3) \\ T_\lambda^x &\stackrel{d}{\leq} (c - x)(c - a) \exp \left\{ \lambda \max_{[x, c]} W - \lambda \min_{[a, c]} W \right\} L' \leq (c - a)^2 L' e^{\lambda D}, \\ L' &= \max_{y \leq 1} L(-\infty, 1, y), \end{aligned}$$

we have

$$\begin{aligned} &P \left\{ T_\lambda^x > e^{\lambda(D+\delta)} \right\} \\ &\leq P \left\{ (c - a)^2 L' e^{\lambda D} > e^{\lambda(D+\delta)} \right\} \\ &= P \left\{ L' > e^{\lambda \delta} / (c - a)^2 \right\} \rightarrow 0, \quad \lambda \rightarrow \infty. \end{aligned}$$

To obtain an estimate from below first we notice that

$$(1.3) \quad \lim_{\lambda \rightarrow \infty} \lambda^{-1} \log C_\lambda = D,$$

---


$$3) \quad \max_I W = \max\{W(x), x \in I\}, \quad \min_I W = \min\{W(x), x \in I\}.$$

where

$$C_\lambda = |\widehat{S}_\lambda(a)| \wedge |\widehat{S}_\lambda(c)| ,$$

and the convergence is uniform in  $x \in [u, v]$ . Next, for given  $\delta > 0$  we set

$$a_1 = \sup\{x < b_1 : W(x) = \delta/4\} ,$$

$$\widehat{s}_\lambda(y) = \widehat{S}_\lambda(y)/C_\lambda ,$$

$$L_\lambda = \min\{L(-1, 1, y) : \widehat{s}_\lambda(a_1) \leq y \leq \widehat{s}_\lambda(b_1)\} .$$

Then applying Lemma 1 and 2 we have

$$\begin{aligned} T_\lambda^x &\stackrel{d}{=} C_\lambda \int_a^c L(\widehat{s}_\lambda(a), \widehat{s}_\lambda(c), \widehat{s}_\lambda(y)) e^{-\lambda W(y)} dy \\ &\geq C_\lambda \int_{a_1}^{b_1} L(-1, 1, \widehat{s}_\lambda(y)) e^{-\lambda W(y)} dy \\ &\geq e^{\lambda(D-\frac{\delta}{4})} (b_1 - a_1) L_\lambda \exp\left\{-\lambda \max_{[a_1, b_1]} W\right\} \\ &= (b_1 - a_1) L_\lambda e^{\lambda(D-\frac{\delta}{2})} . \end{aligned}$$

Since  $\lambda^{-1} \log |\widehat{s}_\lambda(a_1)|$  and  $\lambda^{-1} \log |\widehat{s}_\lambda(b_1)|$  converges to  $\max_{[x \wedge a_1, x \vee a_1]} W - D$ ,

$\max_{[x \wedge b_1, x \vee b_1]} W - D$ , respectively, which are both negative, we have

$$\lim_{\lambda \rightarrow \infty} \widehat{s}_\lambda(a_1) = \lim_{\lambda \rightarrow \infty} \widehat{s}_\lambda(b_1) = 0 ,$$

the convergence being uniform in  $x \in [u, v]$ . Therefore

$$P\left\{T_\lambda^x < e^{\lambda(D-\delta)}\right\} \leq P\left\{L_\lambda < (b_1 - a_1)^{-1} e^{-\lambda\delta/2}\right\} \rightarrow 0 , \quad \lambda \rightarrow \infty$$

uniformly in  $x \in [u, v]$ , because  $\lim_{\lambda \rightarrow \infty} L_\lambda = L(-1, 1, 0) > 0$ .

## §2. The limit distribution of $X(e^{\lambda r}, \lambda W)$

In this section we change the notation slightly. Given  $W \in \mathcal{W}$  and a valley  $V = (a, b_1, b_2, c)$  of  $W$ , we set

$$\Omega = C([0, \infty) \rightarrow \mathbb{R}) ,$$

$$\widehat{\Omega} = C([0, \infty) \rightarrow [a, c]) ,$$

and denote by  $P_\lambda^x$ ,  $x \in \mathbb{R}$  (resp.  $\widehat{P}_\lambda^y$ ,  $y \in [a, c]$ ) the probability measure

on  $\Omega$  (resp.  $\hat{\Omega}$ ) induced by the diffusion process with generator

$$(2.1) \quad \frac{1}{2} e^{\lambda W(x)} \frac{d}{dx} (e^{-\lambda W(x)} \frac{d}{dx})$$

(resp. the diffusion process on  $[a, c]$  with (local) generator (2.1) and with reflecting barriers at  $a$  and  $c$ ). The latter diffusion has the invariant probability measure  $m_\lambda$  given by

$$m_\lambda(dy) = e^{-\lambda W(y)} dy / \int_a^c e^{-\lambda W(z)} dz .$$

For any interval  $[u, v] \subset [a, c]$

$$m_\lambda([u, v]) = \frac{\int_0^\infty e^{-\lambda \xi} K([u, v], \xi) d\xi}{\int_0^\infty e^{-\lambda \xi} K([a, c], \xi) d\xi}$$

where, for an interval  $I$  in  $\mathbb{R}$ ,  $K(I, \xi)$  is the local time at  $\xi$  for the reflected Brownian medium, i.e.,

$$(2.2) \quad K(I, \xi) = \lim_{\varepsilon \downarrow 0} \frac{1}{\varepsilon} \int_I \pi_{[\xi, \xi+\varepsilon)}(W(s)) ds .$$

Therefore

$$(2.3) \quad m_\lambda([u, v]) = \frac{\int_0^\infty e^{-\xi} K([u, v], \lambda^{-1} \xi) d\xi}{\int_0^\infty e^{-\xi} K([a, c], \lambda^{-1} \xi) d\xi} \\ \rightarrow \frac{K([u, v], 0)}{K([a, c], 0)} \equiv m([u, v]) , \quad \lambda \rightarrow \infty .$$

Next we set

$$\hat{P}_\lambda = \int_a^b m_\lambda(dy) \hat{P}_\lambda^y , \quad \mathbb{P}_\lambda^{x,y} = P_\lambda^x \otimes \hat{P}_\lambda^y , \quad \mathbb{P}_\lambda^x = P_\lambda^x \otimes \hat{P}_\lambda .$$

$$R = R(\omega, \hat{\omega}) = \inf\{t \geq 0 : \omega(t) = \hat{\omega}(t)\} .$$

Lemma 5. For any  $\delta > 0$

$$\lim_{\lambda \rightarrow \infty} P_\lambda^0 \left\{ R < e^{\lambda(A+\delta)} \right\} = 1 .$$

Proof. First we prove that

$$(2.4) \quad \lim_{\lambda \rightarrow \infty} \mathbb{P}_\lambda^x \left\{ R < e^{\lambda(A+\delta)} \right\} = 1 \quad \text{holds for } x = b_1 \text{ and } b_2 .$$

Without loss of generality we may consider the case  $x = b_2$ . We write  $b$  instead of  $b_2$  for simplicity. For any  $\delta > 0$  such that  $A + \delta < D$  we define  $a_1 \in (a, b_1)$ ,  $a_2 \in (a, b_1)$ ,  $c_2 \in (b_2, c)$  by

$$\begin{aligned} a_1 &= \max \left\{ x < b_1 : W(x) = A + \frac{\delta}{4} \right\}, \\ a_2 &= \max \left\{ x < b_1 : W(x) = A + \frac{\delta}{2} \right\}, \\ c_2 &= \min \left\{ x > b_2 : W(x) = A + \frac{\delta}{2} \right\}, \end{aligned}$$

and set

$$\begin{aligned} T_0 &= T_0(\omega) = \inf \left\{ t \geq 0 : w(t) = a_1 \right\}, \\ T_1 &= T_1(\omega) = \inf \left\{ t \geq 0 : w(t) \notin (a_1, c_2) \right\}, \\ T_2 &= T_2(\omega) = \inf \left\{ t \geq 0 : w(t) \notin (a_2, c_2) \right\}. \end{aligned}$$

Then we can prove easily that

$$(2.5) \quad \mathbb{P}_\lambda^b \{T_0 < \infty\} \geq \mathbb{P}_\lambda^b \{T_0 = T_1\} = \frac{S_\lambda(c_2) - S_\lambda(b)}{S_\lambda(c_2) - S_\lambda(a_1)} \rightarrow 1, \lambda \rightarrow \infty,$$

and hence

$$\begin{aligned} (2.6) \quad & \mathbb{P}_\lambda^b \{R \leq T_0\} \\ & \geq \mathbb{P}_\lambda^b \{ \hat{w}(0) \in [a, b], \hat{w}(T_0) \in [a_1, c] \} \\ & \geq \mathbb{P}_\lambda^b \{ \hat{w}(0) \in [a, b] \} + \mathbb{P}_\lambda^b \{ \hat{w}(T_0) \in [a_1, c] \} - 1 \\ & = m_\lambda([a, b]) + \int_0^\infty \hat{\mathbb{P}}_\lambda \{ \hat{w}(t) \in [a_1, c] \} \mathbb{P}_\lambda^b \{T_0 \in dt\} - 1 \\ & \rightarrow 1, \lambda \rightarrow \infty, \end{aligned}$$

by (2.3) because  $m(\{x \in (a, c) : W(x) = 0\}) = 1$ . On the other hand Lemma 4 applied to the valley  $(a_2, b_1, b_2, c_2)$  whose depth is  $A + (\delta/2)$  implies

$$(2.7) \quad \mathbb{P}_\lambda^b \{T_1 < e^{\lambda(A+\delta)}\} \geq \mathbb{P}_\lambda^b \{T_2 < e^{\lambda(A+\delta)}\} \rightarrow 1, \lambda \rightarrow \infty,$$

and so

$$\begin{aligned} & \mathbb{P}_\lambda^x \{R < e^{\lambda(A+\delta)}\} \\ & \geq \mathbb{P}_\lambda^x \{T_0 < e^{\lambda(A+\delta)}\} - o(1) \quad (\text{by (2.6)}) \\ & \geq \mathbb{P}_\lambda^x \{T_1 < e^{\lambda(A+\delta)}, T_1 = T_0\} - o(1) \end{aligned}$$



$$\begin{aligned} &\geq P_{\lambda}^x \{T_1 < e^{\lambda(A+\delta)}\} - o(1) && \text{(by (2.5))} \\ &\rightarrow 1, \quad \text{as } \lambda \rightarrow \infty && \text{(by (2.7))}. \end{aligned}$$

Next, to consider the case where the diffusion starts at 0 we shall consider three diffusion processes starting at 0,  $b_1$  and  $b_2$ , respectively. By making use of the comparison theorem in one-dimensional diffusion processes (for example, see [3: p.352]) we can construct, on a suitable probability space  $(\tilde{\Omega}_{\lambda}, \tilde{P}_{\lambda})$ , three processes  $\tilde{X}_0(t)$ ,  $\tilde{X}_1(t)$  and  $\tilde{X}_2(t)$  such that the probability measure on  $\Omega$  induced by  $\tilde{X}_0(t)$  (resp.  $\tilde{X}_1(t)$ ,  $\tilde{X}_2(t)$ ) coincides with  $P_{\lambda}^0$  (resp.  $P_{\lambda}^{b_1}$ ,  $P_{\lambda}^{b_2}$ ) and

$$(2.8) \quad \tilde{X}_1(t) \leq \tilde{X}_0(t) \leq \tilde{X}_2(t), \quad \forall t \geq 0, \quad \tilde{P}_{\lambda}\text{-a.s.}$$

Put

$$\begin{aligned} \tilde{P}_{\lambda} &= \tilde{P}_{\lambda} \otimes \hat{P}_{\lambda}, \\ \tilde{R}_i &= \inf\{t \geq 0 : \tilde{X}_i(t) = \hat{\omega}(t)\}, \quad i = 0, 1, 2. \end{aligned}$$

Since  $\tilde{R}_0 \leq \tilde{R}_1 \vee \tilde{R}_2$  by (2.8), we have

$$\begin{aligned} P_{\lambda}^0 \{R < e^{\lambda(A+\delta)}\} &= \tilde{P}_{\lambda} \{\tilde{R}_0 < e^{\lambda(A+\delta)}\} \\ &\geq \tilde{P}_{\lambda} \{\tilde{R}_1 \vee \tilde{R}_2 < e^{\lambda(A+\delta)}\} \\ &\geq P_{\lambda}^{b_1} \{R < e^{\lambda(A+\delta)}\} + P_{\lambda}^{b_2} \{R < e^{\lambda(A+\delta)}\} - 1 \\ &\rightarrow 1, \quad \lambda \rightarrow \infty \end{aligned}$$

by (2.4), completing the proof of Lemma 5.

**Lemma 6.** For any  $r_1, r_2$  with  $A < r_1 < r_2 < D$  and for any interval  $[u, v]$  in  $\mathbb{R}$

$$\lim_{\lambda \rightarrow \infty} P_{\lambda}^0 \{\omega(e^{\lambda r}) \in [u, v]\} = m([u, v] \cap [b_1, b_2])$$

uniformly in  $r \in [r_1, r_2]$ , where  $m$  is defined in (2.3).

**Proof.** Denote by  $T$  (resp.  $\hat{T}$ ) the exit time of  $(a, c)$  for  $\omega(t)$  (resp.  $\hat{\omega}(t)$ ), and by  $\tilde{T}_R$  (resp.  $\hat{T}_R$ ) the exit time of  $(a, c)$  for  $\omega(t)$  (resp.  $\hat{\omega}(t)$ ) after the collision time  $R$ . Since  $m_{\lambda}(U) \rightarrow 1$  as  $\lambda \rightarrow \infty$  for any open set  $U$  containing  $\{x \in (a, c) : W(x) = 0\}$ , it follows from Lemma 4 that

$$\hat{P}_{\lambda} \{e^{\lambda(D-\delta)} < \hat{T} < e^{\lambda(D+\delta)}\}$$

$$\begin{aligned}
&= \int_a^c m_\lambda(dx) P_\lambda^x \left\{ e^{\lambda(D-\delta)} < T < e^{\lambda(D+\delta)} \right\} \\
&\rightarrow 1, \lambda \rightarrow \infty.
\end{aligned}$$

This combined with Lemma 5 implies

$$\begin{aligned}
p_\lambda &:= \mathbb{P}_\lambda^0 \left\{ R < e^{\lambda r_1} < e^{\lambda r_2} < \hat{T}_R \right\} \\
&\geq \mathbb{P}_\lambda^0 \left\{ R < e^{\lambda r_1} < e^{\lambda r_2} < \hat{T} \right\} \quad (\because \hat{T} \leq \hat{T}_R) \\
&\rightarrow 1, \lambda \rightarrow \infty.
\end{aligned}$$

Therefore for  $r \in [r_1, r_2]$

$$\begin{aligned}
(2.9) \quad &P_\lambda^0 \left\{ \omega(e^{\lambda r}) \in [u, v] \right\} \\
&\geq \mathbb{P}_\lambda^0 \left\{ R < e^{\lambda r_1}, \omega(e^{\lambda r}) \in [u, v], e^{\lambda r_2} < \tilde{T}_R \right\} \\
&= \mathbb{P}_\lambda^0 \left\{ R < e^{\lambda r_1}, \hat{\omega}(e^{\lambda r}) \in [u, v], e^{\lambda r_2} < \hat{T}_R \right\} \\
&\geq p_\lambda + m_\lambda([u, v]) - 1 \\
&\rightarrow m([u, v] \cap [b_1, b_2]), \lambda \rightarrow \infty;
\end{aligned}$$

as for the above equality we used the strong Markov property. Similarly we have

$$\lim_{\lambda \rightarrow \infty} P_\lambda^0 \left\{ \omega(e^{\lambda r}) \in [u, v]^c \right\} \geq m([u, v]^c \cap [b_1, b_2]),$$

which combined with (2.9) implies

$$P_\lambda^0 \left\{ \omega(e^{\lambda r}) \in [u, v] \right\} \rightarrow m([u, v] \cap [b_1, b_2]), \lambda \rightarrow \infty.$$

The uniform convergence in  $r \in [r_1, r_2]$  is also clear.

### §3. Proof of Theorem 1

Let  $V = (a, b_1, b_2, c)$  be a valley of  $W$  such that  $A < 1 < D$ . Such a valley exists with  $Q$ -probability 1. In fact,  $b_1$  and  $b_2$  are taken as

$$b_1 = \text{the smallest root of } W(x) = 0 \text{ in } (a', 0)$$

$$b_2 = \text{the largest root of } W(x) = 0 \text{ in } (0, c')$$

where  $a' = \sup\{x < 0 : W(x) = 1\}$  and  $c' = \inf\{x > 0 : W(x) = 1\}$ . The endpoints  $a$  and  $c$  can be chosen suitably so that  $a < a'$ ,  $c > c'$  and

$V = (a, b_1, b_2, c)$  is a valley with  $A < 1 < D$ . In what follows  $V = (a, b_1, b_2, c)$  denotes such a valley of  $W$ . We denote by  $m_W$  the probability measure on  $\mathbb{R}$  defined by

$$(3.1) \quad m_W([u, v]) = \frac{K([u', v'], 0)}{K([b_1, b_2], 0)}$$

where  $[u', v'] = [u, v] \cap [b_1, b_2]$ . Then, in the notation of §1 Lemma 6 reads as follows: For any interval  $I$  in  $\mathbb{R}$  and for any family  $\{r(\lambda), \lambda > 0\}$  satisfying  $\lim_{\lambda \rightarrow \infty} r(\lambda) = 1$ ,

$$(3.2) \quad \lim_{\lambda \rightarrow \infty} P\{X(e^{\lambda r(\lambda)}), \lambda W \in I\} = m_W(I)$$

for almost all  $W$  with respect to  $Q$ . Now we define  $P = P \otimes Q$  and  $\mu = \int m_W Q(dW)$ . Integrating both sides of (3.2) with respect to  $Q$  we have

$$(3.3) \quad \lim_{\lambda \rightarrow \infty} P\{X(e^{\lambda r(\lambda)}), \lambda W \in I\} = \mu(I).$$

Next, define  $W_\lambda$  as in Lemma 3. Then  $\{W_\lambda(x), x \in \mathbb{R}\}$  is again a reflected Brownian medium. Therefore (3.3) yields

$$(3.4) \quad \lim_{\lambda \rightarrow \infty} P\{X(e^{\lambda r(\lambda)}), \lambda W_\lambda \in I\} = \mu(I).$$

We now apply the scaling relation (1.2) to (3.4); the result is

$$\lim_{\lambda \rightarrow \infty} P\{\lambda^{-2} X(\lambda^4 e^{\lambda r(\lambda)}), W \in I\} = \mu(I).$$

Taking  $r(\lambda) = 1 - 4\lambda^{-1} \cdot \log \lambda$  in the above, we obtain

$$\lim_{\lambda \rightarrow \infty} P\{\lambda^{-2} X(e^\lambda), W \in I\} = \mu(I).$$

This completes the proof of Theorem 1.

#### §4. Proof of Theorem 2

The absolute continuity of  $\mu$  can be proved easily. In fact, if  $\mu_n$  is the measure in  $\mathbb{R}$  defined by

$$\mu_n(I) = E^Q \left\{ \frac{K(I \cap [b_1, b_2])}{K([b_1, b_2])} ; K([b_1, b_2]) > \frac{1}{n} \right\},$$

then  $\mu_n$  is absolutely continuous because

$$\begin{aligned} \mu_n(I) &\leq n E^Q \{ K(I \cap [b_1, b_2]) \} \\ &= 2n \int_I p(|x|, 0, 0) dx, \end{aligned}$$

where  $p(t, \xi, \eta)$  is the transition density of the Brownian motion with absorbing barriers at  $\pm 1$ . Thus  $\mu$  is absolutely continuous because  $\mu_n \uparrow \mu$  as  $n \uparrow \infty$ .

We proceed to the proof of (3). Let  $K(I) = K(I, 0)$  be the local time at 0 for the reflected Brownian medium as defined by (2.2) with  $\xi = 0$  and consider the number of times  $d_\varepsilon(t)$  that the reflected Brownian path  $\{W(u) : u \geq 0\}$  crosses down from  $\xi > 0$  to 0 before time  $t$ . Then as found in [4: p.48]

$$(4.1) \quad Q \left\{ \lim_{\varepsilon \downarrow 0} 2\varepsilon d_\varepsilon(t) = K([0, t]), t \geq 0 \right\} = 1.$$

Let  $a'$ ,  $c'$ ,  $b_1$  and  $b_2$  be defined as in the beginning of §3.

Lemma 7. For  $\alpha, \beta > 0$

$$(4.2) \quad E^Q \left\{ e^{-\alpha K([0, b_2]) - \beta c'} \right\} = \frac{1}{2\alpha + c(\beta)} \cdot \frac{2\sqrt{2\beta}}{e^{\sqrt{2\beta}} - e^{-\sqrt{2\beta}}},$$

where

$$c(\beta) = \frac{e^{\sqrt{2\beta}} + e^{-\sqrt{2\beta}}}{e^{\sqrt{2\beta}} - e^{-\sqrt{2\beta}}} \cdot \sqrt{2\beta}.$$

In Particular,  $K([0, b_2])$  is exponentially distributed:

$$(4.3) \quad E^Q \left\{ e^{-\alpha K([0, b_2])} \right\} = \frac{1}{2\alpha + 1}.$$

Proof. Since  $c(\beta) \sim 1$  as  $\beta \downarrow 0$ , (4.3) follows from (4.2) by letting  $\beta \downarrow 0$ . To prove (4.2) we first apply (4.1) to write down

$$\begin{aligned} (4.4) \quad & E^Q \left\{ e^{-\alpha K([0, b_2]) - \beta c'} \right\} \\ &= E^Q \left\{ e^{-\alpha K([0, c']) - \beta c'} \right\} \\ &= \lim_{\varepsilon \downarrow 0} E^Q \left\{ e^{-2\alpha \varepsilon d_\varepsilon(c') - \beta c'} \right\} \\ &= \lim_{\varepsilon \downarrow 0} \sum_{n=0}^{\infty} e^{-2\alpha \varepsilon n} E^Q \left\{ e^{-\beta T_\varepsilon} \right\}^{n+1} E^Q \left\{ e^{-\beta T_0}; T_0 < T_1 \right\}^n E^Q \left\{ e^{-\beta T_1}; T_1 < T_0 \right\}, \end{aligned}$$

where  $E_\varepsilon^Q$  denotes the expectation with respect to the probability measure of the reflected Brownian motion starting at  $\varepsilon$  and

$$T_x = \inf \{ u \geq 0 : W(u) = x \}.$$

If we set

$$A_{\varepsilon} = e^{-2\alpha\varepsilon} E^Q \left\{ e^{-\beta T_{\varepsilon}} \right\} E_{\varepsilon}^Q \left\{ e^{-\beta T_0}; T_0 < T_1 \right\},$$

$$B_{\varepsilon} = E^Q \left\{ e^{-\beta T_{\varepsilon}} \right\} E_{\varepsilon}^Q \left\{ e^{-\beta T_1}; T_1 < T_0 \right\},$$

then (4.4) yields

$$(4.5) \quad E^Q \left\{ e^{-\alpha K([0, b_2]) - \beta c'} \right\} = \lim_{\varepsilon \downarrow 0} B_{\varepsilon} \sum_{n=0}^{\infty} A_{\varepsilon}^n$$

$$= \lim_{\varepsilon \downarrow 0} \frac{B_{\varepsilon}}{1 - A_{\varepsilon}}.$$

Next we make use of the well-known formula

$$E_x \left\{ e^{-\alpha T_a}; T_a < T_b \right\} = \frac{e^{-\sqrt{2\alpha}(b-x)} - e^{-\sqrt{2\alpha}(b-a)}}{e^{-\sqrt{2\alpha}(b-a)} - e^{-\sqrt{2\alpha}(b-a)}}, \quad a \leq x \leq b,$$

where  $E_x$  denotes the expectation with respect to the probability measure of a standard Brownian motion starting at  $x$ . We then have

$$(4.6) \quad E^Q \left\{ e^{-\beta T_{\varepsilon}} \right\} = 2E_0 \left\{ e^{-\beta T_{\varepsilon}}; T_{\varepsilon} < T_{-\varepsilon} \right\}$$

$$= \frac{2(e^{\varepsilon\sqrt{2\beta}} - e^{-\varepsilon\sqrt{2\beta}})}{e^{2\varepsilon\sqrt{2\beta}} - e^{-2\varepsilon\sqrt{2\beta}}}$$

$$= 1 + o(\varepsilon^2), \quad \varepsilon \downarrow 0;$$

$$(4.7) \quad E_{\varepsilon}^Q \left\{ e^{-\beta T_0}; T_0 < T_1 \right\} = \frac{e^{\sqrt{2\beta}(1-\varepsilon)} - e^{-\sqrt{2\beta}(1-\varepsilon)}}{e^{\sqrt{2\beta}} - e^{-\sqrt{2\beta}}}$$

$$= 1 - \frac{\sqrt{2\beta}(e^{\sqrt{2\beta}} + e^{-\sqrt{2\beta}})}{e^{\sqrt{2\beta}} - e^{-\sqrt{2\beta}}} \cdot \varepsilon + o(\varepsilon), \quad \varepsilon \downarrow 0;$$

$$(4.8) \quad E_{\varepsilon}^Q \left\{ e^{-\beta T_1}; T_1 < T_0 \right\} = \frac{e^{\sqrt{2\beta}\varepsilon} - e^{-\sqrt{2\beta}\varepsilon}}{e^{\sqrt{2\beta}} - e^{-\sqrt{2\beta}}}$$

$$\sim \frac{2\sqrt{2\beta}}{e^{\sqrt{2\beta}} - e^{-\sqrt{2\beta}}} \cdot \varepsilon, \quad \varepsilon \downarrow 0.$$

From (4.6), (4.7) and (4.8) we obtain

$$\frac{B_\varepsilon}{1 - A_\varepsilon} \sim \frac{1}{2\alpha + c(\beta)} \cdot \frac{2\sqrt{2\beta}}{e^{\sqrt{2\beta}} - e^{-\sqrt{2\beta}}} , \quad \varepsilon \downarrow 0 ,$$

which combined with (4.5) proves the lemma.

Given  $x > 0$  we set

$$K_1 = K([b_1, 0]), \quad K_2 = K([0, x]), \quad K_3 = K([x, b_2]) .$$

Lemma 8. For  $x > 0$  and  $t > 0$

$$(4.9) \quad \begin{aligned} & E^Q \left\{ K_3 e^{-t(K_1 + K_2 + K_3)} ; x < b_2 \right\} \\ &= \frac{2}{(2t + 1)^3} E^Q \left\{ (1 - W(x)) e^{-tK([0, x])} ; x < c' \right\} . \end{aligned}$$

Proof. The left hand side of (4.9) equals

$$E^Q \left\{ e^{-tK_1} \right\} E^Q \left\{ K_3 e^{-t(K_2 + K_3)} ; x < b_2 \right\} .$$

Since  $E^Q \left\{ e^{-tK_1} \right\} = (2t + 1)^{-1}$  by Lemma 7, for the proof of the lemma it is enough to show

$$(4.10) \quad \begin{aligned} & E^Q \left\{ K_3 e^{-t(K_2 + K_3)} ; x < b_2 \right\} \\ &= \frac{2}{(2t + 1)^2} E^Q \left\{ (1 - W(x)) e^{-tK_2} ; x < c' \right\} . \end{aligned}$$

To prove this we introduce the smallest  $\sigma$ -field  $\mathcal{F}_x$  on  $W$  which makes  $W(s)$ ,  $0 \leq s \leq x$ , measurable and consider the event  $\Gamma$  that the shifted trajectory  $W(\cdot + x)$  hits 0 before it hits 1. Then first using the strong Markov property of the reflected Brownian motion and then (4.3), we have

$$\begin{aligned} & E^Q \left\{ K_3 e^{-tK_3} \mathbb{1}_\Gamma / \mathcal{F}_x \right\} \\ &= \{1 - W(x)\} E^Q \left\{ K([0, b_2]) e^{-tK([0, b_2])} \right\} \\ &= \frac{2}{(2t + 1)^2} \{1 - W(x)\} , \quad \text{a.s.} \end{aligned}$$

Since  $\{x < b_2\} = \{x < c'\} \cap \Gamma$  and  $\{x < c'\} \in \mathcal{F}_x$ , we have

$$\begin{aligned} & E^Q \left\{ K_3 e^{-t(K_2 + K_3)} ; x < b_2 \right\} \\ &= E^Q \left\{ e^{-tK_2} \mathbb{1}_{\{x < c'\}} E^Q \left\{ K_3 e^{-tK_3} \mathbb{1}_\Gamma / \mathcal{F}_x \right\} \right\} \end{aligned}$$

$$= \frac{2}{(2t+1)^2} E^Q \left\{ (1 - W(x)) e^{-tK_2}; x < c' \right\},$$

proving (4.10) and hence the lemma.

Lemma 9. For  $\lambda > 0$  and  $t > 0$

$$(4.11) \quad \int_0^\infty e^{-\lambda x} E^Q \left\{ (1 - W(x)) e^{-tK([0, x])}; x < c' \right\} dx \\ = \frac{1}{\lambda} \left\{ 1 - \frac{(2t+1)S}{C + 2tS} \right\},$$

where

$$C = \cosh \sqrt{2\lambda}, \quad S = \frac{\sinh \sqrt{2\lambda}}{\sqrt{2\lambda}}.$$

Proof. Let  $\varphi(x) = 1 - |x|$ . Consulting with [4: Chapter 5], we see that the left hand side of (4.11) equals  $f_\lambda(0)$  where  $f_\lambda$  is the continuous solution of

$$(4.12) \quad \begin{cases} \lambda f - \frac{1}{2}f'' = \varphi & \text{in } (-1, 0) \cup (0, 1) \\ \frac{1}{2}\{f'(0+) - f'(0-)\} = 2tf(0) \\ f(-1) = f(1) = 0. \end{cases}$$

To solve (4.12) we first find the solution  $g_\lambda$  of  $\lambda f - \frac{1}{2}f'' = \varphi$  in  $(-1, 1)$  with boundary condition  $f(-1) = f(1) = 0$  and then express  $f_\lambda$  as follows:

$$f_\lambda(x) = \begin{cases} g_\lambda(x) + c \sinh \sqrt{2\lambda}(1+x) & \text{for } x \in (-1, 0) \\ g_\lambda(x) + c \sinh \sqrt{2\lambda}(1-x) & \text{for } x \in (0, 1) \end{cases}.$$

If we determine  $c$  so that the above  $f_\lambda$  satisfies the second condition of (4.12), then the  $f_\lambda$  is a solution of (4.12). Thus  $f_\lambda(0)$  can be computed and we obtain (4.11).

Now Theorem 2 can be proved as follows. By Lemma 8 we have

$$\mu((x, \infty)) = E^Q \left\{ \frac{K((x, x \vee b_2])}{K([b_1, b_2])} \right\} \\ = E^Q \left\{ \frac{K_3}{K_1 + K_2 + K_3}; x < b_2 \right\} \\ = \int_0^\infty E^Q \left\{ K_3 e^{-t(K_1 + K_2 + K_3)}; x < b_2 \right\} dt$$

$$= \int_0^{\infty} \frac{2}{(2t+1)^3} E^Q \left\{ (1 - W(x)) e^{-tK([0,x])} ; x < c \right\} dt$$

and hence by Lemma 9

$$\begin{aligned} \int_0^{\infty} e^{-\lambda x} \mu((x, \infty)) dx &= \int_0^{\infty} \frac{2}{(2t+1)^3} \cdot \frac{1}{\lambda} \left\{ 1 - \frac{(2t+1)S}{c+2tS} \right\} dt \\ &= \frac{1}{2\lambda} - \frac{1}{\lambda} \int_0^{\infty} \frac{2}{(2t+1)^2} \cdot \frac{S}{c+2tS} dt. \end{aligned}$$

Thus integration by parts yields

$$\begin{aligned} \int_0^{\infty} e^{-\lambda x} \mu(dx) &= \frac{1}{2} - \lambda \int_0^{\infty} e^{-\lambda x} \mu((x, \infty)) dx \quad (\text{notice that } \mu((0, \infty)) = \frac{1}{2}) \\ &= \int_0^{\infty} \frac{2S}{(2t+1)^2(c+2tS)} dt \\ &= \int_0^{\infty} \frac{Sdt}{(t+1)^2(c+tS)}, \end{aligned}$$

and this proves (3).

#### REFERENCES

- [1] T. Brox, A one-dimensional diffusion process in a Wiener medium, to appear in Ann. Probab.
- [2] A.O. Golosov, The limit distributions for random walks in random environments, Soviet Math. Dokl., 28(1983), 18-22.
- [3] N. Ikeda and S. Watanabe, Stochastic Differential Equations and Diffusion Processes, North-Holland, 1981.
- [4] K. Itô and H.P. McKean, Diffusion Processes and Their Sample Paths, Springer-Verlag, 1965.
- [5] K. Kawazu, Y. Tamura and H. Tanaka, One-dimensional diffusions and random walks in random environments, in preparation.
- [6] H. Kesten, The limit distribution of Sinai's random walk in random environment, to appear in Physica.
- [7] S. Schumacher, Diffusions with random coefficients, Contemporary Math. (Particle Systems, Random Media and Large Deviations, ed. by R. Durrett), 41(1985), 351-356.



- [8] Y.G.Sinai, The limiting behavior of a one-dimensional random walk in a random medium, Theory of Probab. and its Appl. 27(1982), 256-268.
- [9] H.Tanaka, Limit distributions for one-dimensional diffusion processes in self-similar random environments, to appear in the Proc. of the workshop on Hydrodynamic behavior and interacting particle systems and applications held at IMA, University of Minnesota, March 1986.

Department of Mathematics  
Faculty of Science and Technology  
Keio University  
Yokohama, Japan