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THE FIRST PASSAGE PROBLEM FOR GENERALIZED ORNSTEIN-UHLENBECK

PROCESSES WITH NON-POSITIVE JUMPS \*

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1. Introduction. Let  $(\Omega, \mathcal{F}, P)$  be a probability space. We consider a cadlag stationary random process  $S_t, t \geq 0$ , with independent increments and non-positive jumps

$\Delta S_t = S_t - S_{t-} = S_t - \lim_{s \uparrow t} S_s \leq 0$ , that is defined on this space and satisfies  $S_0 = 0$ .

It is well known ([3]) that the characteristic function of  $S_t$  has the form

$$(1.1) \quad E \exp(iuS_t) = \exp t \{ibu - cu^2 + \int_{(-\infty, 0)} F(dx) (e^{iux} - 1 - iux \cdot 1_{\{x \geq -1\}})\},$$

where  $-\infty < b < \infty$ ,  $c \geq 0$ , and the Lévy measure  $F(\cdot)$  satisfies

$$(1.2) \quad \int_{(-\infty, 0)} F(dx) |x|^2 < \infty.$$

Following Skorokhod ([8]) one can use the analytical continuation of (1.1) to the half-plane  $\text{Re}(iu) > 0$  and obtain the Laplace transform of  $S_t$  by substituting

$u$  instead of  $iu$ . Thus, we have

$$(1.3) \quad E \exp(uS_t) = \exp t\psi(u), \quad u \geq 0,$$

where

$$(1.4) \quad \psi(u) = bu + cu^2 + \int_{(-\infty, 0)} F(dx) (e^{ux} - 1 - ux \cdot 1_{\{x \geq -1\}}).$$

For arbitrary  $\lambda > 0$  and  $-\infty < x < \infty$  we define the random process  $X_t, t \geq 0$ , by the formula

$$(1.5) \quad X_t = e^{-\lambda t} \left( x + \int_{(0, t]} e^{\lambda v} dS_v \right),$$

the stochastic integral w.r.t. the semi-martingale  $S$  being understood in the usual sense.

Definition. The random process  $X$  will be called the starting at  $x$  generalized

Ornstein-Uhlenbeck process with parameter  $\lambda > 0$ .

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Certainly, the process  $X$  is characterized by the triplet  $(b, c, F(\cdot))$  as well. With  $b = 0$ ,  $c = 1/2$  and  $F(\cdot) = 0$  our definition yields the standard Wiener process  $S$  and the usual Ornstein-Uhlenbeck process  $X$ .

Given a real number  $\mu > x$ , let us introduce the first passage time

$$(1.6) \quad T_\mu(x) = \inf \{t \geq 0 : X_t \geq \mu\}.$$

As far as  $\Delta X_t = \Delta S_t \leq 0$ , if  $T_\mu(x) < \infty$  one gets immediately the equality

$$X_{T_\mu(x)} = \mu.$$

The purpose of this paper is to determine the distribution of  $T_\mu(x)$ ,  $\mu > x$ , by means of Laplace transform

$$(1.7) \quad \gamma_\mu(\theta, x) = E \exp(-\theta T_\mu(x)), \quad \theta > 0.$$

It should be noted that generally speaking, we have no equation for the transition density of  $X$  and the usual Darling-Siebert approach to the first passage problem of diffusion processes ([2]) is not applicable in our case. Our approach is based on martingale techniques and depends essentially on the existence of suitable martingales on the process  $X$  (see Theorem 1 below). Besides the new generality of the explicit representation for  $\gamma_\mu(\theta, x)$  (Section 4), this approach gives us in particular the possibility to obtain ones again and in a natural way the interesting result of Novikov ([6]) concerning the first passage times of a stable process  $S$  through one-sided non-linear boundaries. The basic tool in this special case is the suitable time-change (Section 6) that transfers the linear problems for  $X_t$ ,  $t \geq 0$ , into some non-linear problems for  $S_t$ ,  $t \geq 0$ , and conversely. We make use of the reconversion in order to give an example of optimal stopping problem that admits a solution in terms of  $T_\mu(x)$ .

2. The process  $X$ . For the next we need to calculate the conditional Laplace transforms of the process  $X$  that was defined in (1.5). Let us introduce the  $\sigma$ -algebras

$$F_t^X = \sigma(X_s, 0 \leq s \leq t); \quad t \geq 0, \quad \text{and the functions } L(u; t, s) = E\{\exp(uX_t) | F_s^X\}, s < t, u > 0.$$

Since the stochastic integral in (1.5) might be looked at as an integral taken in the sense of convergence in probability ([4]), a simple argument leads to the following result.

Proposition 1. For any  $0 \leq s < t$  and  $u \geq 0$  one has

$$(2.1) \quad L(u; t, s) = \exp\left\{e^{-\lambda(t-s)} X_s \cdot u + \int_s^t \psi(u \cdot e^{-\lambda(t-v)}) dv\right\}.$$

Proof. With an arbitrary subdivision  $s = t_0 < t_1 < \dots < t_n = t$ ,  $\varepsilon = \max_{i \leq n} |t_i - t_{i-1}|$

and  $Y_t = \int_{(0, t]} e^{\lambda v} dS_v$ , we get

$$\begin{aligned} E\{\exp(uY_t) | F_s^X\} &= \exp(uY_s) \cdot E\left\{\exp\left(u \int_{(s, t]} e^{\lambda v} dS_v\right) | F_s^X\right\} \\ &= \exp(uY_s) \lim_{\varepsilon \rightarrow 0} \prod_{i=1}^n E \exp(u \cdot e^{\lambda t_{i-1}} \cdot (S_{t_i} - S_{t_{i-1}})) \\ &= \exp(uY_s) \lim_{\varepsilon \rightarrow 0} \prod_{i=1}^n \exp\{\psi(u \cdot e^{\lambda t_{i-1}})(t_i - t_{i-1})\} \\ &= \exp(uY_s + \int_s^t \psi(u \cdot e^{\lambda v}) dv) \end{aligned}$$

as a consequence of (1.3) and the independent increments property of  $S$ .

Now starting with (1.5) we have

$$\begin{aligned} L(u; t, s) &= \exp(e^{-\lambda t} \cdot xu) E\{\exp(u \cdot e^{-\lambda t} \cdot Y_t) | F_s^X\} \\ &= \exp\{e^{-\lambda t} \cdot xu + e^{-\lambda t} \cdot Y_s u + \int_s^t \psi(u \cdot e^{-\lambda(t-v)}) dv\} \end{aligned}$$

and the latter obviously implies (2.1).

Corollary 1. The Laplace transform of  $X_t$  has the form

$$E \exp(uX_t) = \exp\left\{e^{-\lambda t} \cdot xu + \int_0^t \psi(u \cdot e^{-\lambda(t-v)}) dv\right\}, u \geq 0.$$

Corollary 2. The process  $X$  is a cadlag Markov process. (Certainly,  $X$  has also the strong Markov property.)

3. The martingale  $M$ . We are going to introduce a martingale  $M_t(\theta)$ ,  $t \geq 0$ , depending on the process  $X$  trajectories. To this end, one observes that because of (1.2) the quantity  $F[-1, -z]$  is finite for every  $z$ ,  $0 < z \leq 1$ . Thus, the measure

$$G(dz) = F[-1, -z] dz$$

on  $(0, 1]$  is well defined. We need the following assumption.

Hypothesis G. Either  $c > 0$  or the measure  $G(\cdot)$  satisfies the condition

$$(3.1) \quad \lim_{z \rightarrow 0^+} z^\kappa \cdot G(z, 1] = C > 0$$

for some constant  $\kappa$ ,  $0 < \kappa < 1$ .

Next, one defines successively

$$(3.2) \quad g(y) = -\frac{1}{\lambda} \int_1^y \frac{\psi(u)}{u} du, \quad y > 0,$$

and

$$(3.3) \quad M_t(\theta) = e^{-\theta t} \cdot \int_0^\infty y^{\frac{\theta}{\lambda} - 1} \cdot \exp\{X_t \cdot y + g(y)\} dy, \quad t \geq 0.$$

The next statement is crucial because it permits an essential use of the martingale theory later on.

Theorem 1. Under the hypothesis G for any positive  $\theta$  the random process  $M_t(\theta)$ ,  $t \geq 0$ , is a martingale w.r.t.  $F_t^X$ ,  $t \geq 0$ .

Proof. First, we observe that our hypothesis G implies the convergence of the integral in (3.3). In fact, we have

$$g(y) = -\frac{b}{\lambda} (y - 1) - \frac{c}{2\lambda} (y^2 - 1) - \frac{1}{\lambda} g_1(y) - \frac{1}{\lambda} g_2(y),$$

where

$$g_1(y) = \int_1^y \frac{\psi_1(u)}{u} du, \quad g_2(y) = \int_1^y \frac{\psi_2(u)}{u} du$$

and

$$\psi_1(u) = \int_{(-\infty, -1)} F(dx) (e^{ux} - 1), \quad \psi_2(u) = \int_{[-1, 0)} F(dx) (e^{ux} - 1 - ux), \quad u \geq 0.$$

The convergence of the integral at  $y = 0$  is obvious, because  $\lim_{y \rightarrow 0^+} g(y) \geq -\infty$ .

Now let us denote  $d_1 = \int_{(-\infty, -1)} F(dx) \geq 0$ ,  $d_2 = \int_{[-1, 0)} F(dx) x^2 \geq 0$ . In consequence

of (1.2) one gets  $0 \leq d_1 + d_2 < \infty$ . Our function  $\psi_1$  satisfies  $0 \geq \psi_1(u) + d_1$  and  $0 \geq \frac{\psi_1(u)}{u} + 0$  as  $u \rightarrow \infty$ . This means that  $|g_1(y)| \leq \int_1^y \frac{|\psi_1(u)|}{u} du \leq d_1 \ln y$ . On the

other hand  $0 < e^{ux} - 1 - ux \leq \frac{u x^2}{2}$ ,  $u > 0$ ,  $-1 \leq x < 0$ , and in this way one

obtains the inequalities  $0 \leq \frac{\psi_2(u)}{u} \leq \frac{u}{2} \cdot d_2 < \infty$  and  $0 \leq g_2(y) \leq \frac{d_2}{4} (y^2 - 1)$ .

If  $c > 0$ , the corresponding term  $-\frac{c}{2\lambda}(y^2 - 1)$  in  $g(y)$  ensures the convergence. If  $c = 0$ , by the equality  $\frac{\psi_2(u)}{u} = \frac{1}{0} \int (1 - e^{-uz}) G(dz)$ , where obviously  $0 \leq \int_0^1 z G(dz) = \frac{d_2}{2} < \infty$ , the hypothesis (3.1) and the corollary of Theorem 4.15 in [1] one gets  $\lim_{u \rightarrow \infty} u^{-\kappa} \cdot \frac{\psi_2(u)}{u} \geq C \cdot \Gamma(1 - \kappa) > 0$ . Consequently,  $\frac{\psi_2(u)}{u} \geq C_2 \cdot u^\kappa$  for any  $C_2$  belonging to the interval  $(0, C \cdot \Gamma(1 - \kappa))$  and  $u \geq u_2(C_2) > 0$  (sufficiently large). This implies  $g_2(y) \geq C_2 \cdot y^{1+\kappa} + C_1$ ,  $y > u_2(C_2)$ , and the convergence of our integral too.

Secondly, applying Fubini's lemma and (2.1) for  $0 \leq s \leq t$  (and with  $z = ye^{-\lambda(t-s)}$ ) we get

$$\begin{aligned} E\{M_t(\theta) \mid F_S^X\} &= e^{-\theta t} \cdot \int_0^\infty \frac{\theta}{y\lambda}^{-1} E\{\exp(X_t \cdot y + g(y)) \mid F_S^X\} dy \\ &= e^{-\theta s} \cdot \int_0^\infty \frac{\theta}{y\lambda}^{-1} \exp\{g(y) - \theta(t-s) + e^{-\lambda(t-s)} y \cdot X_s + \int_s^t \psi(ye^{-\lambda(t-v)}) dv\} dy \\ &= e^{-\theta s} \cdot \int_0^\infty \frac{\theta}{z\lambda}^{-1} \exp\{zX_s + g(ze^{\lambda(t-s)}) + \int_0^{t-s} \psi(ze^{\lambda v}) dv\} dz. \end{aligned}$$

But the function  $f(u, z) = g(ze^{\lambda u}) + \int_0^u \psi(ze^{\lambda v}) dv$ ,  $u \geq 0$ , satisfies the condition

$$\frac{\partial f(u, z)}{\partial u} = g'(ze^{\lambda u}) \cdot z\lambda e^{\lambda u} + \psi(ze^{\lambda u}) = g'(y) \cdot \lambda y + \psi(y) \equiv 0$$

with  $y = ze^{\lambda u}$ , in view of (3.2). Therefore,

$$f(u, z) = \text{const} = f(0, z) = g(z)$$

and we get  $E\{M_t(\theta) \mid F_S^X\} = X_s$ , that completes the proof.

Remark 1. We emphasize the fact that Theorem 1 is valid for every process  $X$  with  $S$  containing a Gaussian component ( $c > 0$ ). If the process  $S$  has no Gaussian component ( $c = 0$ ), the condition (3.1) is nevertheless fulfilled for a class of measures  $F(\cdot)$  that includes the stable processes  $S$  with parameter  $\alpha$  satisfying  $1 < \alpha < 2$ . Because of its importance, we consider this special case in Section 5.

4. The Laplace transform of  $T_\mu(x)$ . Now we are in a position to derive an explicit expression for the Laplace transform  $\gamma_\mu(\theta, x)$ . Due to the particular structure of

the martingale  $M(\theta)$  we have the following result.

Theorem 2. Under the hypothesis G the next equality holds:

$$(4.1) \quad \gamma_{\mu}(\theta, x) = \frac{\int_0^{\infty} y^{\frac{\theta}{\lambda}-1} \exp(xy + g(y)) dy}{\int_0^{\infty} y^{\frac{\theta}{\lambda}-1} \exp(\mu y + g(y)) dy}, \quad \theta > 0.$$

Proof. We put  $T_{\mu}(x)\Delta t$  instead of  $t$  in (3.3) and we make use of the well known

martingale property that

$$E M_{T_{\mu}(x)\Delta t}(\theta) = E M_0(\theta) = \int_0^{\infty} y^{\frac{\theta}{\lambda}-1} \exp(xy + g(y)) dy.$$

Next, one observes that

$$0 \leq M_{T_{\mu}(x)\Delta t}(\theta) \leq \int_0^{\infty} y^{\frac{\theta}{\lambda}-1} \exp(\mu y + g(y)) dy$$

and, moreover, when  $T_{\mu}(x) = \infty$  then

$$0 \leq M_{T_{\mu}(x)\Delta t}(\theta) = M_t(\theta) \leq e^{-\theta t} \int_0^{\infty} y^{\frac{\theta}{\lambda}-1} \exp(\mu y + g(y)) dy$$

as well. Therefore,

$$\lim_{t \rightarrow \infty} E M_{T_{\mu}(x)\Delta t}(\theta) = E M_{T_{\mu}(x)} \cdot 1_{\{T_{\mu}(x) < \infty\}} = \int_0^{\infty} y^{\frac{\theta}{\lambda}-1} \exp(\mu y + g(y)) dy \cdot \gamma_{\mu}(\theta, x).$$

The right-hand sides of our equalities give directly (4.1).

Remark 2. For the validity of Theorem 2 we need not (and we did not use) any fact

about the finiteness of  $T_{\mu}(x)$ . It is well known that  $T_{\mu}(x) < \infty$  P-a.s. if and only

if  $\lim_{\theta \rightarrow 0} \gamma_{\mu}(\theta, x) = 1$ . The latter equality is easily verified when there exists

$\lim_{y \rightarrow 0} g(y) > -\infty$  or when  $\int_{(-\infty, -1)} F(dx) |x| < \infty$ .

5. The case of stable process S with parameter  $1 < \alpha \leq 2$ . Now we turn to the par-

ticular case when the following hypothesis is satisfied.

Hypothesis  $H_{\alpha}$ . Either  $F(\cdot) = 0$  and  $c > 0$  (we characterize this by posing  $\alpha = 2$ ),

or  $c = 0$  and  $F(dx) = \frac{\sigma \cdot dx}{|x|^{\alpha+1}} 1_{\{x < 0\}}$  for some  $\sigma > 0$  and  $1 < \alpha < 2$ .

Using standard arguments (see [8], §25, Theorem 4) one obtains the equivalent form of  $H_{\alpha}$  in the terms of our function  $\psi$ :  $H_{\alpha}$ ,  $1 < \alpha \leq 2$ , means that

$$(5.1) \quad \psi(u) = \overline{\psi}(u) = \overline{b}u + \overline{\sigma}u^{\alpha}$$

with some  $\bar{b}$ ,  $-\infty < \bar{b} < \infty$ , and  $\bar{\sigma} > 0$ . In this situation by (3.2) we get

$$(5.2) \quad g(y) = \bar{g}(y) = -\frac{\bar{b}}{\lambda} (y - 1) - \frac{\bar{\sigma}}{\alpha\lambda} (y^\alpha - 1),$$

and the martingale  $M(\theta)$  is well defined via (3.3).

Following Novikov we introduce the function

$$H(\nu, \alpha, x) = \frac{1}{\Gamma(-\alpha\nu)} \int_0^\infty y^{-\alpha\nu - 1} \exp(xy - \frac{1}{\alpha} y^\alpha) dy,$$

which turns to be analytic in the half-plane  $\operatorname{Re} \nu < 1$ . All the essential properties of  $H(\nu, \alpha, x)$  are collected in the supplement of [6].

Next we obtain a special case of Theorem 2.

Proposition 2. Under the hypothesis  $H_\alpha$ ,  $1 < \alpha \leq 2$ , the following equality holds

for  $\theta > 0$ :

$$(5.3) \quad \gamma_\mu(\theta, x) = \frac{H\left(-\frac{\theta}{\alpha\lambda}, \alpha, \left(\frac{\lambda}{\bar{\sigma}}\right)^\alpha \left(x - \frac{\bar{b}}{\lambda}\right)\right)}{H\left(-\frac{\theta}{\alpha\lambda}, \alpha, \left(\frac{\lambda}{\bar{\sigma}}\right)^\alpha \left(\mu - \frac{\bar{b}}{\lambda}\right)\right)}.$$

Moreover, this formula defines also an analytical continuation of the Laplace transform  $\gamma_\mu(\theta, x)$  to the half-plane  $\operatorname{Re} \theta > -\alpha\lambda \cdot \nu_\alpha(\bar{\mu})$ , where  $\bar{\mu} = \left(\frac{\lambda}{\bar{\sigma}}\right)^\alpha \left(\mu - \frac{\bar{b}}{\lambda}\right)$  and  $\nu_\alpha(z)$  is the smallest positive zero of  $H(\nu, \alpha, z)$  with  $(\alpha, z)$  fixed.

Proof. Applying the change of variables  $y = \left(\frac{\lambda}{\bar{\sigma}}\right)^\alpha z$  we see the formula (5.3) is another form of (4.1) for  $\theta > 0$ . As far as the right-hand side of (5.3) is analytic in  $\theta$  in the half-plane  $\operatorname{Re} \theta > -\alpha\lambda \cdot \nu_\alpha(\bar{\mu})$  (see [6]), the left-hand side can be analytically continued in  $\theta$  to this half-plane.

Corollary 3. Since  $\lim_{\nu \rightarrow 0} H(\nu, \alpha, x) = 1$ ,  $-\infty < x < \infty$ , under the hypothesis  $H_\alpha$  we get

$$\lim_{\theta \downarrow 0} \gamma_\mu(\theta, x) = 1 \text{ and, consequently, } T_\mu(x) < \infty \quad \text{P-a.s.}$$

6. The time change - two applications. Throughout this section we suppose the hypothesis  $H_\alpha$  holds with some  $\alpha$ ,  $1 < \alpha \leq 2$ , and  $\bar{b} = 0$  (see (5.1)). As a consequence we have

$$\psi(u) = \bar{\psi}(u) = \bar{\sigma} \cdot u^\alpha, \quad 1 < \alpha \leq 2,$$

and the process  $X$  is stationary too (see (2.1)).

Let us introduce the real (increasing and continuous) function

$$\delta(t) = (\alpha\lambda)^{-1} (e^{\alpha\lambda t} - 1), \quad t \geq 0,$$

which determines an one-to-one mapping of  $[0, \infty)$  onto  $[0, \infty)$ , and the convers function

$$\rho(t) = (\alpha\lambda)^{-1} \ln(1 + \alpha\lambda t), \quad t \geq 0.$$

Lemma 1. The distributions of  $S_t$ ,  $t \geq 0$ , and of  $\tilde{S}_t = \int_0^{\rho(t)} e^{\lambda v} dS_v$ ,  $t \geq 0$ , coincide.

Proof. As in Proposition 1 one calculates

$$E \exp(u\tilde{S}_t) = E \exp(uY_{\rho(t)}) = \exp\{\bar{\sigma}u^\alpha \cdot \delta(\rho(t))\} = \exp(\bar{\sigma}u^\alpha t), \quad u > 0.$$

But under the hypothesis stated ( $H_\alpha$  and  $\bar{b} = 0$ ) the latter term is just  $E \exp(uS_t)$ .

The lemma is proved.

Now for any constants  $a$ ,  $b$  and  $c$  such that  $b \geq 0$  and  $ab\frac{1}{\alpha} + c > 0$ , define the stopping time  $\tau(a,b,c)$  w.r.t.  $F_t^S$ ,  $t \geq 0$ , by the formula

$$(6.1) \quad \tau(a,b,c) = \inf \{t > 0 : S_t \geq a(t+b)\frac{1}{\alpha} + c\}$$

and pose

$$(6.2) \quad \tau_\mu(x) = \tau(\mu(\alpha\lambda)\frac{1}{\alpha}, (\alpha\lambda)^{-1}, -x), \quad \mu > x.$$

The following simple fact is valid in our situation.

Theorem 3. The stopping time  $T_\mu(x)$  has the same distribution as  $\rho(\tau_\mu(x))$  does.

Proof. We define similarly  $\tilde{\tau}(a,b,c)$  and  $\tilde{\tau}_\mu(x)$  by replacing  $S_t$  by  $\tilde{S}_t$  in (6.1)

and (6.2). Next, starting with (1.6), we calculate

$$\begin{aligned} T_\mu(x) &= \inf \{t : x + Y_t \geq \mu e^{\lambda t}\} \\ &= \inf \{\rho(s) : Y_{\rho(s)} \geq \mu e^{\lambda \rho(s)} - x\} \\ &= \inf \{\rho(s) : \tilde{S}_s \geq \mu(1 + \alpha\lambda s)\frac{1}{\alpha} - x\} = \rho(\tilde{\tau}_\mu(x)). \end{aligned}$$

The statement of the theorem follows from Lemma 1 which says the distribution of  $\tilde{\tau}_\mu(x)$  coincides with the distribution of  $\tau_\mu(x)$ .

From Theorem 3 and Proposition 2 we deduce the following result of A. Novikov

(see [6], Theorem 1).

Theorem 4. For every  $a, b, c$  with  $b \geq 0$ ,  $ab^{\frac{1}{\alpha}} + c > 0$ , one has

$$(6.3) \quad E (\tau(a, b, c) + b)^{\nu} = b^{\nu} \frac{H(\nu, \alpha, -cb^{-\frac{1}{\alpha}}d)}{H(\nu, \alpha, ad)}, \quad \text{if } b > 0 \text{ and } \nu < \nu_{\alpha}(ad),$$

and

$$(6.4) \quad E (\tau(a, b, c)^{\nu}) = \begin{cases} \frac{(cd)^{\alpha\nu}}{H(\nu, \alpha, ad)}, & \text{if } \nu < \nu_{\alpha}(ad), \\ +\infty, & \text{if } \nu \geq \nu_{\alpha}(ad), \end{cases}$$

where  $d = (\alpha c)^{\frac{1}{\alpha}}$ .

Proof. Assume  $b > 0$  and put  $x = -c$ ,  $\lambda = (\alpha b)^{-1}$ ,  $\mu = ab^{\frac{1}{\alpha}}$ . Then

$$\mu - x = ab^{\frac{1}{\alpha}} + c > 0, \quad \bar{\mu} = \left(\frac{\lambda}{\sigma}\right)^{\frac{1}{\alpha}} \mu = ad$$

and by Proposition 2 with  $\nu = -\frac{\theta}{\alpha\lambda}$  we get the equalities

$$\begin{aligned} E (\tau(a, b, c) + b)^{\nu} &= E \left( \bar{\tau}_{\mu}(x) + \frac{1}{\alpha\lambda} \right)^{\nu} \\ &= b^{\nu} \cdot E (\alpha\lambda \bar{\tau}_{\mu}(x) + 1)^{\nu} = b^{\nu} \cdot E \exp\{\nu \ln(1 + \alpha\lambda \bar{\tau}_{\mu}(x))\} \\ &= b^{\nu} \cdot E \exp\{-\theta \rho(\bar{\tau}_{\mu}(x))\} = b^{\nu} \cdot \frac{H(\nu, \alpha, -cb^{-\frac{1}{\alpha}}d)}{H(\nu, \alpha, ad)}, \end{aligned}$$

provided that  $\theta > -\alpha\lambda\nu_{\alpha}(ad)$  (or  $\nu < \nu_{\alpha}(ad)$ ). The rest statements of the theorem

follow from the properties of  $H(\nu, \alpha, x)$ , the case  $b = 0$  being taken into account by letting  $b \rightarrow 0$  (or  $\lambda \rightarrow +\infty$ ).

Remark 3. In the original theorem of Novikov (with  $d = 1$ , see [6]) one makes use of the fact that

$$(t + b)^{\nu} \cdot \frac{H(\nu, \alpha, \frac{S_t - c}{(t + b)^{\frac{1}{\alpha}}})}{(t + b)^{\frac{1}{\alpha}}}, \quad t \geq 0, b > 0,$$

is a complex-valued martingale (w.r.t.  $F_t^S$ ,  $t \geq 0$ ) for every complex  $\nu$  with  $\text{Re } \nu < 1$ .

This fact involves an analytical continuation in contrast to our Theorem 1.

As a second example we consider an optimal stopping problem originally treated in more general setting in [5], [7] and [9]. This problem admits a simple solution in terms of stopping times  $T_{\mu}(x)$ .

Under the hypothesis stated at the beginning of this section ( $H_\alpha$  and  $\bar{b} = 0$ ) the quantity

$$(6.5) \quad v(x, b, \tau) = E \frac{x + S_\tau}{b + \tau}, \quad b > 0, \quad -\infty < x < \infty,$$

is to be maximized on stopping times  $\tau = \tau(\omega)$  w.r.t.  $F_t^S, t \geq 0$ .

By Lemma 1 we have

$$v(x, b, \tau) = v(x, b, \tilde{\tau}) = E \frac{x + \tilde{S}_{\tilde{\tau}}}{b + \tilde{\tau}},$$

using  $\tilde{S}_t = Y_{\rho(t)}, t \geq 0$ , and  $\tilde{\tau}$  in the place of  $S_t, t \geq 0$ , and  $\tau$ . Now taking

$\lambda = \frac{1}{\alpha b}$  and  $t = \delta(s), s \geq 0$ , we get

$$\frac{x + \tilde{S}_t}{b + t} = \frac{x + Y_{\rho(t)}}{b + t} = \frac{e^{\lambda \rho(t)} \cdot X_{\rho(t)}}{\frac{1}{\alpha \lambda} + t} = \frac{e^{\lambda s} \cdot X_s}{\frac{1}{\alpha \lambda} e^{\alpha \lambda s}} = \alpha \lambda e^{-(\alpha - 1)\lambda s} \cdot X_s.$$

Consequently, it is equivalent to consider the problem of maximizing the quantity

$$(6.6) \quad V(x, b, T) = \frac{1}{b} E e^{-\beta T} \cdot X_T, \quad \beta = \frac{\alpha - 1}{\alpha b} > 0,$$

on stopping times  $T = T(\omega)$  w.r.t.  $F_s^X, s \geq 0$ , provided that  $T = \rho(\tau)$ , because

$$V(x, b, T) = v(x, b, \tau).$$

By [7] for  $\alpha = 2$  and [5] for  $1 < \alpha < 2$  one knows the solution of the original problem of maximizing (6.5) is one of the stopping times  $\tau(a, b, -x)$  or the stopping time  $\tau_0 = 0$ .

Let us denote

$$\Psi(\mu) = \frac{\int_0^\infty y^{\alpha-2} \exp(\mu y - \bar{\sigma} b y^\alpha) dy}{\int_0^\infty y^{\alpha-1} \exp(\mu y - \bar{\sigma} b y^\alpha) dy}, \quad -\infty < \mu < \infty.$$

As far as  $\Psi(\mu)$  is positive, decreasing and continuous and  $\Psi(0) = \Gamma(\frac{\alpha-1}{\alpha}) > 0$ ,

the equation  $\mu = \Psi(\mu)$  has a unique solution  $\tilde{\mu}$  (moreover,  $0 < \tilde{\mu} < \Psi(0)$ ). The

corresponding result in our case is given below without proof because it can be justified as in [5] and [7] (see also [9], Example 2, for the case  $\alpha = 2$  and  $\lambda = 1$ ).

**Theorem 5.** For every real  $x$  and  $b > 0$ , either the stopping time  $T_{\tilde{\mu}}(x)$ , or the stopping time  $T_x(x) = 0$  maximizes the quantity (6.6). More precisely,

$$\sup_T V(x,b,T) = V(x,b,T_{\tilde{\mu}}) = \frac{\tilde{\mu}}{b} \gamma_{\tilde{\mu}}(\beta, x) \quad \text{if } x \leq \tilde{\mu},$$

and

$$\sup_T V(x,b,T) = V(x,b,0) = \frac{x}{b} \quad \text{if } x > \tilde{\mu}.$$

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