# A balanced excited random walk 

## Une marche excité équilibrée

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## A R T I C L E I N F O

## Article history:

Received 7 September 2010
Accepted after revision 21 February 2011
Available online 10 March 2011
Presented by Marc Yor


#### Abstract

The following random process on $\mathbb{Z}^{4}$ is studied. At first visit to a site, the two first coordinates perform a (2-dimensional) simple random walk step. At further visits, it is the last two coordinates which perform a simple random walk step. We prove that this process is almost surely transient. The lower dimensional versions are discussed and various generalizations and related questions are proposed.


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## R É S U M É

Nous étudions le processus suivant sur $\mathbb{Z}^{4}$. À la première visite en un site, les deux premières coordonnées effectuent un saut d'une marche simple (2-dimensionnelle). Aux visites suivantes en ce site, ce sont les deux dernières coordonnées qui effectuent un saut de marche simple. Nous montrons que ce processus est presque sûrement transitoire. Nous discutons également des dimensions inférieures et divers généralisations et questions connexes sont proposées.
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## 1. Introduction

The excited random walk as defined by Benjamini and Wilson [2] has a bias in some fixed direction, a feature which is highly useful in its analysis. See e.g. [10] and references within. Attempts to relax the dependence of the proof structure on monotonicity resulted in a number of works where the walker has competing drifts. See [1,6,4]. One motivation was to get closer to standard models of reinforced random walks on $\mathbb{Z}^{d}$, which are symmetric in nature. We think for instance on the question of recurrence vs. transience of 1 -reinforced random walks, which is still widely open (see the surveys [11] and [12]). With the same goal in mind, we started exploring excited-like models where the walker is in addition also a martingale or a bounded perturbation of one, and posed some questions in 2007 [7], which, it seems, are all still open. Progress on this kind of models was achieved in [5], but there the laws were not nearest-neighbors. Here we describe and solve one such model of a nearest-neighbor walk in 4 dimensions.

We describe a general form of the model in any dimension $d \geqslant 2$, but we will actually only deal with 4 dimensions here. So one has first to choose arbitrarily two integers $d_{1} \geqslant 1$ and $d_{2} \geqslant 1$ such that $d=d_{1}+d_{2}$. Then we define the process $\left(S_{n}, n \geqslant 0\right)$ on $\mathbb{Z}^{d}$ as a mixture of two simple random walks in the following sense. Set $S_{n}=\left(X_{n}, Y_{n}\right)$, where $X_{n} \in \mathbb{Z}^{d_{1}}$ is the set of the first $d_{1}$ coordinates of $S_{n}$ and $Y_{n} \in \mathbb{Z}^{d_{2}}$ is the set of the last $d_{2}$ coordinates. Now the rule is the following. First $S_{0}=0$. Next if $S$ visits a site for the first time then only the $X$ component performs a simple random walk step, that is:

[^0]\[

\mathbb{P}\left[S_{n+1}-S_{n}=(0, ···, 0, \pm 1,0, ···, 0) \mid \mathcal{F}_{n}\right]= $$
\begin{cases}1 /\left(2 d_{1}\right) & S_{n} \neq S_{i} \forall i<n \\ 0 & \text { otherwise }\end{cases}
$$
\]

where the $\pm 1$ can be at any of the first $d_{1}$ coordinates. Otherwise, only $Y$ performs a simple random walk step:

$$
\mathbb{P}\left[S_{n+1}-S_{n}=(0, \ldots, 0, \pm 1,0, \ldots, 0) \mid \mathcal{F}_{n}\right]= \begin{cases}1 /\left(2 d_{2}\right) & \exists i<n S_{i}=S_{n} \\ 0 & \text { otherwise }\end{cases}
$$

where this time the $\pm 1$ can be at any of the last $d_{2}$ coordinates. We call this process $S$ the $M\left(d_{1}, d_{2}\right)$-random walk.
Here we say that a process is transient if almost surely any site is visited only finitely many times. It is said to be recurrent if almost surely it visits all sites infinitely often. We will prove the following:

Theorem 1. The $M(2,2)$-random walk is transient.

The proof of Theorem 1 is elementary, and uses only basic estimates on the standard 2-dimensional simple random walk. It relies on finding good upper bounds for the probability of return to the origin and then uses the Borel-Cantelli Lemma (what makes however the proof nontrivial is that the two components $X$ and $Y$ are not independent).

Note that the canonical projections of $S$ on $\mathbb{Z}^{d_{1}}$ and $\mathbb{Z}^{d_{2}}$ are usual (time changed) simple random walks. So if $d=4$, and if $d_{1}$ or $d_{2}$ equals 3 , then $S$ is automatically transient, since the simple random walk on $\mathbb{Z}^{3}$ is transient. Likewise if $d$ is larger than 5 , then for any choice of $d_{1}$ and $d_{2}$, the resulting process will be transient. Thus the question of recurrence vs. transience is only interesting in dimension less than or equal to 4 . In dimension 3 there are two versions: $d_{1}=1$ and $d_{2}=2$ or $d_{1}=2$ and $d_{2}=1$. We conjecture that in both cases $S$ will be transient, also because it is a 3 -dimensional process. Proving this seems nontrivial, but notice that a possible intermediate step between the dimension 3 and 4 could be to consider the analogue problem on the discrete 3-dimensional Heisenberg group, which is generated by 2 elements (and their inverses), yet balls of radius $r$ have size order $r^{4}$.

Let us make some comments now on the 2-dimensional case. As a 2-dimensional process, we believe that $M(1,1)$ is recurrent. Observe however that this is not true when starting from any configuration of visited sites. Indeed if we start with a vertical line of visited sites, then the process will be trapped in this line, and if the line does not include the origin, the process will not return there. It is also not difficult to construct starting environments such that the first coordinate of the process will tend almost surely toward $+\infty$. For example if the initial configuration is the "trumpet" $\left\{(x, y):|y|<e^{x}\right\}$ then the walker will drift to infinity in the $x$ direction. ${ }^{1}$ Of course it is not possible for the random walk to create these environments in finite time, so it is not an obstacle for recurrence, but it may be interesting to keep this in mind. Another problem concerns the limiting shape of the range (i.e. the set of visited sites) of the process. Based on heuristics and some simulations, we believe that it is a vertical interval. This problem is closely related to the question of evaluating the size of the range $R_{n}$ at time $n$. Indeed the horizontal displacement of the process at time $n$ is of order $\sqrt{\# R_{n}}$, whereas its vertical displacement is always of order $\sqrt{n}$. So another formulation of the problem would be to show that $R_{n}$ is sublinear. By the way we mention a related question. Assume that at each step, one can decide, conditionally on the past, to move the first coordinate or the second coordinate (and then perform a 1-dimensional simple random walk step). Then what is the best strategy to maximize the range? In particular is it possible for the range to be of size roughly $n$, or at least significantly larger than $n / \ln n$, which is the size of the range of the simple random walk?

A possible generalization of our model would be to consider multi-excited versions, in the spirit of Zerner [13]. In this case one should first decompose $d$ as $d=d_{1}+\cdots+d_{m}$, for some $m \geqslant 2$ and $d_{i} \geqslant 1$, with $i \leqslant m$. Then at $i$ th visit to a site only the $i$ th component of $S$ performs a simple random walk step, if $i<m$, and at further visits only the $m$ th component moves. In dimension 4 for instance the case $d_{1}=2$ and $d_{2}=d_{3}=1$ seems interesting and nontrivial. Another interesting case is $d \geqslant 3$ and $d_{i}=1$ for each $i \leqslant d$ (even the case $d$ very large seems nontrivial).

A related problem which appeared in draft versions of this paper was solved by Y. Peres and P. Sousi (private communication) who proved, among other things, the following. Let $\mu_{1}$ and $\mu_{2}$ be any two symmetric laws on $\mathbb{Z}^{d}, d \geqslant 3$, and assume the support of $\mu_{1}$ and $\mu_{2}$ both generate $\mathbb{Z}^{d}$. Decide that at first visit to a site the jump of the process has law $\mu_{1}$, and at further visits it has law $\mu_{2}$. Then this process is transient.

## 2. Proof of the theorem

The theorem is a direct consequence of the following proposition:

Proposition 1. There exists a constant $C>0$ such that for any $n>1$,

$$
\mathbb{P}\left[0 \in\left\{S_{n}, \ldots, S_{2 n}\right\}\right] \leqslant C\left(\frac{\ln \ln n}{\ln n}\right)^{2}
$$

[^1]Indeed assuming this proposition we get

$$
\sum_{k \geqslant 0} \mathbb{P}\left[0 \in\left\{S_{2^{k}}, \ldots, S_{2^{k+1}}\right\}\right]<+\infty
$$

and we can conclude by using the Borel-Cantelli Lemma. So all we have to do is to prove this proposition.
Proof of Proposition 1. For any $n \geqslant 1$, denote by $r_{n}$ the cardinality of the range of $S$ at time $n$. The next lemma will be needed:

Lemma 1. For any $M>0$, there exists a constant $C>0$, such that

$$
\mathbb{P}\left[n /(C \ln n)^{2} \leqslant r_{n} \leqslant 99 n / 100\right]=1-o\left(n^{-M}\right)
$$

Proof. Note first that for any $k$, if $S_{k}$ and $S_{k+1}$ were not already visited in the past, then $S_{k+2}=S_{k}$ with probability at least $1 / 4$. In particular for any $k$, there is probability at least $1 / 4$ for the event
$E_{k}:=\{S$ is not at a fresh site at one of the times $k, k+1$ or $k+2\}$,
to occur. In other words, if $\varepsilon_{k}=1_{E_{k}}$, then $\mathbb{E}\left[\varepsilon_{3 k} \mid \varepsilon_{0}, \ldots, \varepsilon_{3(k-1)}\right] \geqslant 1 / 4$ for all $k$. Then a standard use of the Azuma-Hoeffding inequality gives that

$$
\mathbb{P}\left[\sum_{k=0}^{[n / 3]} \varepsilon_{3 k} \leqslant n / 100\right]=o\left(n^{-M}\right),
$$

which gives the desired upper bound on $r_{n}$.
We now prove the lower bound. Let $c>0$ be fixed. Let $\left(U_{n}, n \geqslant 0\right)$ be a simple random walk on $\mathbb{Z}^{2}$. For any $n \geqslant 1$ and $x \in \mathbb{Z}^{2}$, denote by $N_{n}(x)$ the number of visits of $U$ to $x$ before time $n$. A simple and standard calculation (see e.g. [9, Proposition 4.2.4]) shows that there exists a constant $C>0$ such that the probability to not visit $x$ in the next $n$ steps after a given visit is $\geqslant C / \log n$. Using the strong Markov property one gets that the probability to make $k+1$ visits by time $n$ is $\leqslant \exp (-C k / \log n)$ and hence there exists some $C^{\prime}>0$ depending on $M$ such that

$$
\mathbb{P}\left[N_{n}(x) \geqslant C^{\prime}(\ln n)^{2}\right]=o\left(n^{-M-2}\right)
$$

Moreover since $U$ makes nearest-neighbor jumps, before time $n$ it stays in a ball of radius $n$. Thus if $N_{n}^{*}=\sup _{x} N_{n}(x)$, then

$$
\mathbb{P}\left[N_{n}^{*} \geqslant C^{\prime}(\ln n)^{2}\right]=n^{2} \times o\left(n^{-M-2}\right)=o\left(n^{-M}\right)
$$

Thus if $r_{n, U}$ is the size of the range of $U$ at time $n$, we get

$$
\mathbb{P}\left[r_{n, U} \leqslant n /\left(C^{\prime}(\ln n)^{2}\right)\right]=o\left(n^{-M}\right)
$$

Let's come back to the original process $S=(X, Y)$ now. We just observe that at time $n$ one of the $X$ or $Y$ component performed $n / 2$ steps. Since each of these components is a simple random walk, we deduce from the previous estimate, that before time $n, X$ or $Y$ will visit at least $n /\left(2 C^{\prime}(\ln n)^{2}\right)$ sites, with probability at least $1-o\left(n^{-M}\right)$. This gives the desired lower bound for $r_{n}$ and concludes the proof of the lemma.

We can finish now the proof of Proposition 1. As noticed in the introduction, observe that the $X$ and $Y$ components are time changed simple random walks. Specifically we have the equality in law:

$$
\begin{equation*}
\left(\left(X_{k}, Y_{k}\right), k \geqslant 1\right)=\left(\left(U\left(r_{k-1}\right), V\left(k-r_{k-1}\right)\right), k \geqslant 1\right) \tag{1}
\end{equation*}
$$

where $U$ and $V$ are two independent simple random walks on $\mathbb{Z}^{2}$, and where by abuse of notation we also denote by $r_{k}$ the size of the range of the $(U, V)$ process at $k$ th step. More precisely $r_{k}$ may be defined recursively by $r_{0}=1$ and for $k \geqslant 1$,

$$
r_{k}:=\#\left\{(U(0), V(0)), \ldots,\left(U\left(r_{k-1}\right), V\left(k-r_{k-1}\right)\right)\right\} .
$$

The proof of (1) is immediate by induction.
By using now Lemma 1 and the independence of $U$ and $V$, we get

$$
\mathbb{P}\left[0 \in\left\{S_{n}, \ldots, S_{2 n}\right\}\right] \leqslant \mathbb{P}\left[0 \in\left\{U\left(n /(C \ln n)^{2}\right), \ldots, U(2 n)\right\}\right] \mathbb{P}[0 \in\{V(n / 100), \ldots, V(2 n)\}]+o\left(n^{-M}\right)
$$

Thus Proposition 1 follows from the following lemma:
Lemma 2. Let $U$ be the simple random walk on $\mathbb{Z}^{2}$ and let $t \in\left[n /(\ln n)^{3}, 2 n\right]$. Then

$$
\begin{equation*}
\mathbb{P}[0 \in\{U(t), \ldots, U(2 n)\}]=\mathcal{O}\left(\frac{\ln \ln n}{\ln n}\right) \tag{2}
\end{equation*}
$$

Proof. This lemma is standard, but we give a proof for the reader's convenience. First let $|\cdot|$ denote the $L^{1}$ norm on $\mathbb{R}^{2}$ (i.e. $\left|\left(x_{1}, x_{2}\right)\right|=\left|x_{1}\right|+\left|x_{2}\right|$ for all $\left.\left(x_{1}, x_{2}\right) \in \mathbb{R}^{2}\right)$. Since $t \geqslant n /(\ln n)^{3}$, it is well known (see e.g. [3, Theorem 1, p. 542]) that

$$
\mathbb{P}\left[|U(t)| \leqslant \frac{\sqrt{n}}{(\ln n)^{3}}\right]=\mathcal{O}\left((\ln n)^{-1}\right)
$$

Moreover it is known (see e.g. [8, Proposition 1.6.7]) that

$$
\mathbb{P}_{x}\left[\tau_{0}<\tau_{|x|(\ln |x|)^{4}}\right]=\mathcal{O}\left(\frac{\ln \ln |x|}{\ln |x|}\right),
$$

where $\mathbb{P}_{x}$ denotes the law of $U$ starting from $x$ and for any $r \geqslant 0$,

$$
\tau_{r}=\inf \{k>0: r<|U(k)| \leqslant r+1\} .
$$

But if $|x| \geqslant \sqrt{n} /(\ln n)^{3}$, then $\sqrt{n} \ln n=\mathcal{O}\left(|x|(\ln |x|)^{4}\right)$. On the other hand (see e.g. [9, Proposition 2.1.2])

$$
\mathbb{P}\left[\tau_{\sqrt{n} \ln n} \leqslant 2 n\right]=\mathcal{O}\left(n^{-1}\right)
$$

Notice finally that if $|x| \geqslant \sqrt{n} /(\ln n)^{3}$, then

$$
\frac{\ln \ln |x|}{\ln |x|}=\mathcal{O}\left(\frac{\ln \ln n}{\ln n}\right)
$$

The lemma follows by using Markov's property at time $t=n /(\ln n)^{3}$,

$$
\begin{aligned}
\mathbb{P}[0 \in\{U(t), \ldots, U(2 n)\}] & \leqslant \mathbb{P}\left[|U(t)| \leqslant \frac{\sqrt{n}}{(\ln n)^{3}}\right]+\max _{|x|>\sqrt{n} /(\ln n)^{3}} \mathbb{P}_{x}\left(\tau_{0} \leqslant 2 n\right) \\
& \leqslant \mathcal{O}\left((\ln n)^{-1}\right)+\max _{|x|>\sqrt{n} /(\ln n)^{3}} \mathbb{P}_{x}\left[\tau_{0}<\tau_{|x|(\ln |x|)^{4}}\right]+\mathbb{P}\left[\tau_{|x|(\ln |x|)^{4}} \leqslant 2 n\right] \\
& \leqslant \mathcal{O}\left((\ln n)^{-1}\right)+\mathcal{O}\left(\frac{\ln \ln n}{\ln n}\right)+\mathcal{O}\left(n^{-1}\right)=\mathcal{O}\left(\frac{\ln \ln n}{\ln n}\right)
\end{aligned}
$$

as required.
The proof of Theorem 1 is now finished.
Remark 1. A slight modification of the proof actually shows that $M(2,2)$ is transient for any finite initial configuration of visited sites. One just has to replace $r_{k-1}$ in (1) by $r_{k-1}-n_{k-1}$, where $n_{k}$ is the number of sites (different from the origin) which are considered as visited at time 0 , and effectively visited by the ( $U, V$ ) process in the $k$ first steps. Since sup $n_{k}$ is finite this does not change the rest of the proof. This is of course not always the case if this configuration is infinite. For instance if we decide that all sites of the form $(0,0, *, *)$ are already visited at time 0 , then the $X$ component will never move and $M(2,2)$ will not be transient.

## Acknowledgements

This work was done while B.S. was a visitor at the Weizmann Institute and he thanks the institute for its kind hospitality.

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    doi:10.1016/j.crma.2011.02.018

[^1]:    ${ }^{1}$ We will not prove any of these claims, as they are somewhat off-topic.

