CUT POINTS AND DIFFUSIVE RANDOM WALKS
IN RANDOM ENVIRONMENT

POINTS DE COUPURE ET MARCHES ALÉATOIRES
DIFFUSIVES EN MILIEU ALÉATOIRE

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ABSTRACT. – We study in this work a special class of multidimensional random walks in random environment for which we are able to prove in a non-perturbative fashion both a law of large numbers and a functional central limit theorem. As an application we provide new examples of diffusive random walks in random environment. In particular we construct examples of diffusive walks which evolve in an environment for which the static expectation of the drift does not vanish.

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RÉSUMÉ. – On étudie dans cet article une classe de marches aléatoires en milieu aléatoire en dimension supérieure, pour lesquelles on prouve de manière non perturbative une loi des grands nombres et un théorème central limite fonctionnel. Comme application de ces résultats on construit de nouveaux exemples de marches aléatoires diffusives en milieu aléatoire. En particulier on présente des exemples de marches aléatoires diffusives qui évoluent dans un environnement aléatoire pour lequel l’espérance statique de la dérive n’est pas nulle.

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0. Introduction

Over the recent years there has been considerable interest in the study of random walks in random environment. The asymptotic behavior of this canonical model of

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random motion in a random medium remains quite mysterious, especially in the multi-dimensional situation. Recent advances have mainly been concerned with the ballistic situation where the walk has a non-degenerate asymptotic velocity, see [15, 12, 13, 16]. Diffusive behavior has remained largely unexplored, except for the work of Lawler [7] when the walk has no local drift, and of Bricmont and Kupiainen [2], for small isotropic perturbations of the simple random walk in dimension \( d \geq 3 \). The present article provides new examples of walks with diffusive behavior. It studies a special class of walks for which we are able to derive in a non-perturbative fashion the law of large numbers as well as central limit theorems. Interestingly, proofs are simple when compared to [2].

We now describe the setting. We consider two integers \( d_1 \geq 5, d_2 \geq 1 \), and write \( d = d_1 + d_2 \). We view \( \mathbb{Z}^{d_1} \) and \( \mathbb{Z}^{d_2} \) as the respective subspaces of \( \mathbb{Z}^d \) of vectors with vanishing last \( d_2 \) and vanishing first \( d_1 \) components. Throughout this work we study random walks in random environment for which the \( \mathbb{Z}^{d_1} \)-projection evolves according to a standard random walk, and the random environment only affects the \( \mathbb{Z}^{d_2} \)-component. Specifically we consider a number \( \kappa \in (0, \frac{1}{2d}) \) (the ellipticity constant for the \( \mathbb{Z}^{d_1} \)-component) and a \((2d_1 + 1)\)-vector governing the jump-distribution of the \( \mathbb{Z}^{d_1} \)-components of the walk:

\[
(q(e))_{|e| \leq 1, e \in \mathbb{Z}^{d_1}}, \text{ with } \sum_{|e| = 1} q(e) = 1, \quad q(e) = q(-e) > 0, \text{ for } |e| \leq 1, e \in \mathbb{Z}^{d_1},
\]
and

\[
q(e) \geq \kappa, \text{ for } e \neq 0,
\]

and introduce

\[
\mathcal{P}_{q(\cdot)} \text{ the set of } (2d)\text{-vectors } (p(e))_{|e| = 1}, \text{ with } p(e) \in [0, 1], \text{ for all } e \in \mathbb{Z}^d, |e| = 1,
\]

\[
\sum_{|e| = 1} p(e) = 1, \text{ and } p(e) = q(e), \text{ for } e \in \mathbb{Z}^{d_1}, |e| = 1.
\]

The random environment is an element \( \omega = (\omega(x, \cdot))_{x \in \mathbb{Z}^d} \) of \( \Omega = \mathcal{P}_{q(\cdot)}^{\mathbb{Z}^d} \), endowed with the product \( \sigma \)-algebra and the product measure \( \mathbb{P} = \mu \otimes \mathbb{P}_{q(\cdot)}^{\mathbb{Z}^d} \), where \( \mu \) is a probability on \( \mathcal{P}_{q(\cdot)} \) governing the distribution of the environment at a single site. The random walk in the random environment \( \omega \) is the canonical Markov chain \( (X_n)_{n \geq 0} \) on \((\mathbb{Z}^d)^\mathbb{N}\) with state space \( \mathbb{Z}^d \), and “quenched” law \( P_{x, \omega} \) starting from \( x \in \mathbb{Z}^d \), under which

\[
P_{x, \omega}[X_{n+1} = X_n + e \mid X_0, \ldots, X_n] \stackrel{P_{x, \omega}^{\omega \text{-as.}}}{=} \omega(X_n, e), \quad n \geq 0, |e| = 1,
\]

\[
P_{x, \omega}[X_0 = x] = 1.
\]

The annealed laws are then defined as the semi-direct products on \( \Omega \times (\mathbb{Z}^d)^\mathbb{N} \):

\[
P_x = \mathbb{P} \times P_{x, \omega}, \quad \text{for } x \in \mathbb{Z}^d.
\]

Our very choice of environments \( \omega \) in \( \Omega \) forces the \( \mathbb{Z}^{d_1} \)-projection of \( X_n \) to evolve under \( P_{x, \omega} \) as a random walk with jump distribution \( q(\cdot) \). We assume symmetry of \( q(\cdot) \) for otherwise we would be in a non-nestling situation where the law of large numbers and the central limit theorem have been proven in [15, 12]. The assumption \( d_1 \geq 5 \), enables to exploit the presence of cut times of the random walk, where loosely speaking past and future of the random walk have no intersection, (for the precise definition, see (1.4)).
These cut times play a somewhat similar role to the regeneration times employed in [15, 12], although they do not provide a renewal structure.

In the above setting we are able to derive a law of large numbers:

$$ P_0 \text{-a.s., } \frac{X_n}{n} \to v \quad (\text{with a deterministic } v). $$

(0.5)

Further assuming that either the law of the environment is invariant under the antipodal transformation (cf. (2.1), in this case $v = 0$), and $d_1 \geq 7$, or without symmetry assumption that $d_1 \geq 13$, we obtain a functional central limit theorem under the quenched measure:

$$ \mathbb{P}\text{-a.s., under } P_{0,\omega}, \text{ the Skorohod-space valued } B_n = \frac{1}{\sqrt{n}} (X_{\lfloor n \rfloor} - \lfloor n \rfloor v) $$

converges in law to a Brownian motion with deterministic covariance.

(0.6)

One can of course replace the quenched measure by the annealed measure $P_0$ in (0.6).

The above result in particular provides examples of diffusive behavior beyond current knowledge. It can also be applied to certain small perturbations of the standard random walk. For $\varepsilon \in (0, 1)$, following [14], we define

$$ S_\varepsilon = \text{the set of } (2d)\text{-vectors } \{p(e)\}, |e| = 1, \text{ with } |p(e) - \frac{1}{2d}| \leq \frac{\varepsilon}{4d}, \text{ for all } e, $$

and

$$ \sum_e p(e) = 1, \quad (0.7) $$

and write $d(x, \omega)$ for the local drift:

$$ d(x, \omega) = \sum_e \omega(x, e)e. \quad (0.8) $$

It is shown in [14], that for $\eta > 0$, and small $\varepsilon$ depending on $d$ and $\eta$, when the single site distribution is concentrated on $S_\varepsilon$, and the static expectation of the local drift $\mathbb{E}[d(0, \omega)]$ has size bigger than $\varepsilon^{5/2-\eta}$, when $d = 3$, $\varepsilon^{3-\eta}$, when $d \geq 4$, the walk has a non-vanishing limiting velocity (much more is known, see [14]). One can wonder whether the same remains true for arbitrarily small non-vanishing $\mathbb{E}[d(0, \omega)]$. We show here that this is not the case and provide examples when $d \geq 7$, of single site distributions concentrated on $S_\varepsilon$, for arbitrarily small $\varepsilon$, with $\mathbb{E}[d(0, \omega)] \neq 0$, but vanishing limiting velocity $v$, and even with diffusive behavior, when $d \geq 15$. We also construct further examples of analogous behavior for walks which are not small perturbations of the simple random walk.

Let us now explain how this article is organized.

In Section 1, we provide an alternative representation of the law of the walk under the annealed measure which takes advantage of the cut times. We then derive the law of large numbers in Theorem 1.4.

In Section 2, we prove the functional central limit theorem under the annealed measure. The case with antipodal symmetry and $d_1 \geq 7$ is covered by Theorem 2.1, the general case with $d_1 \geq 13$, is treated in Theorem 2.2.

Section 3 explains how the functional central limit theorem under the annealed measure can be strengthened to a similar statement under the quenched measure.
Section 4 provides examples of walks which are small perturbations of the simple random walk, for which \( \mathbb{E}[d(0, \omega)] \neq 0 \), but the limiting velocity vanishes, \( (d \geq 7) \), and which behave diffusively, \( (d \geq 15) \).

Section 5 contains further examples of analogous behavior, which in a certain sense are small perturbations of a one-dimensional random walk in a random environment.

1. An alternative representation of \( P_0 \) and a law of large numbers

In this section we first introduce some further notations and provide a special representation of the walk under the measure \( P_0 \), see Proposition 1.2. This representation will provide an easy comparison of the walk under \( P_0 \) with a process constructed as an additive functional over a probability space with an ergodic shift. This will lead to a law of large numbers, cf. Theorem 1.4.

We begin with some notations. We denote by \((e_i)_{1 \leq i \leq d}\) the canonical basis of \( \mathbb{R}^d \), and by \( | \cdot | \) the Euclidean distance on \( \mathbb{R}^d \). For \( U \) a subset of \( \mathbb{Z}^d \), \( |U| \) denotes the cardinality of \( U \) and \( \partial U \) the boundary of \( U \): \( \partial U = \{ x \in \mathbb{Z}^d \setminus U, \exists y \in U, |x - y| = 1 \} \). The drift will be the \( \mathbb{R}^d \)-valued function on \( \mathbb{P}_q(\cdot) \):

\[
d(p) = \sum_{|e|=1} p(e) e = \sum_{i > d} (p(e_i) - p(-e_i)) e_i, \quad \text{for } p(\cdot) \in \mathbb{P}_q(\cdot).
\]

To represent the random walk governing the evolution of the \( \mathbb{Z}^{d_1} \)-projection of the RWRE, we consider the product space

\[
W_* = \{ e \in \mathbb{Z}^{d_1}, |e| \leq 1 \}^\mathbb{Z},
\]

endowed with the product \( \sigma \)-algebra \( \mathcal{W}_* \) and the product measure \( P = q^{\otimes \mathbb{Z}} \), (in the notation of (0.1)). We denote by \((\theta_n)_{n \in \mathbb{Z}}\) the canonical shift on \( W_* \) and by \((I_n)_{n \in \mathbb{Z}}\) the canonical coordinates. We then define, for \( w \in W_* \),

\[
X^1_n = X^1_n(w) = \begin{cases} 
I_1 + \cdots + I_n, & n \geq 1, \\
0, & n = 0, \\
-(I_{n+1} + \cdots + I_0), & n \leq -1.
\end{cases}
\]

(1.2)

Observe that \( X^1_n, n \geq 0 \), and \( X^1_n, n \leq 0 \), are two independent random walks on \( \mathbb{Z}^{d_1} \) with jump-distribution \( q \), and that

\[
X^1_n \circ \theta_k = X^1_{n+k} - X^1_k, \quad n, k \in \mathbb{Z}.
\]

(1.3)

The set of cut times where “future” and “past” of \( X^1 \) have no intersection will play an important role in this article. Specifically, for \( w \in W_* \), we consider

\[
\mathcal{D}(w) = \{ n \in \mathbb{Z}, X^1_{(-\infty,n-1]}(w) \cap X^1_{[n,\infty)}(w) = \emptyset \},
\]

(1.4)

as well as the stationary point process

\[
N(w, dk) = \sum_{n \in \mathbb{Z}} \delta_n(dk) 1\{ n \in \mathcal{D}(w) \}.
\]

(1.5)
It will be convenient to restrict $P$ to the shift-invariant set of full $P$-measure (cf. Lemma 1.1 below)

$$W = \{ w \in W_*, \ N(w, (-\infty, 0]) = N(w, [0, \infty)) = \infty \}. \quad (1.6)$$

We will write $W$ for the restriction of $W_*$ to $W$. We collect some useful properties relative to the point process $N$ in the following

**Lemma 1.1.**

$$P(0 \in \mathcal{D}) > 0. \quad (1.7)$$

$$P(W) = 1, \quad and \ on \ W, \quad N(w, dk) = \sum_{m \in \mathbb{Z}} \delta_{T_m(w)}(dk), \quad (1.8)$$

where $T_m(w), m \in \mathbb{Z}$ are $\mathbb{Z}$-valued variables on $W$, increasing with $m$ and such that $T_0 \leq 0 < T_1$.

$$\hat{P} \overset{\text{def}}{=} P[\cdot | 0 \in \mathcal{D}] \text{ is invariant under } \hat{\theta} \overset{\text{def}}{=} \theta_{T_1} \cdot \quad (1.9)$$

$$\int T^1 d\hat{P} = P[0 \in \mathcal{D}]^{-1}. \quad (1.10)$$

$$\int f dP = \int \sum_{0}^{T^1-1} f \circ \theta_k d\hat{P} / \int T^1 d\hat{P}, \quad (1.11)$$

for $f$ bounded measurable on $W$.

$$P[ T^1 > n ] \leq c (\log n)^{1+\varepsilon} n^{-\varepsilon/2}, \quad n \geq 1, \quad (1.12)$$

for a positive constant $c$ depending only on $d_1$ and $q(\cdot)$.

**Proof.** The claim (1.7) follows from the fact that $X^1_n, n \geq 0,$ and $X^1_{-n}, n \geq 0,$ are independent random walks on $\mathbb{Z}^{d_1}, d_1 \geq 5,$ with jump distribution $q(\cdot)$ using classical estimates on the decrease of the transition probability, cf. Spitzer [11], p. 75, and similar arguments as in Section 3.2 of Lawler [8] or Section 4 of Erdös and Taylor [5]. Using the ergodicity of $\theta$ and (1.7), $P(W) = 1$ follows and (1.8) is straightforward. Up to a different normalization $P$ is the Palm measure attached to the stationary point process $N$, cf. Neveu [9], Chapter II, (see in particular (10), p. 317). The statements (1.9), (1.10) (Kac’s lemma), and (1.11) are then standard. We now turn to the proof of (1.12).

We consider an integer $L \geq 1$, and write:

$$k_j = 1 + Lj, \quad \text{for } j \geq 0. \quad (1.13)$$

Then for $J \geq 1$:

$$P[ T^1 > k_{2J} ] = P[ N(w, [1, k_{2J}]) = 0] \leq \sum_{0 \leq j < 2J+1} P[ X^1_{(-\infty, k_{j-1}]} \cap X^1_{[k_{j+1}, \infty)} \neq \emptyset]$$

$$+ P[ \text{for all } 0 \leq j < 2J+1, \ X^1_{(-\infty, k_{j-1}]} \cap X^1_{[k_{j+1}, \infty)} = \emptyset, \quad \text{and } N(w, [1, k_{2J}]) = 0]$$

$$\text{for } n \geq 1,$$
We first bound $a_2$. To this end note that when $N(w, [1, k_2J]) = 0$ and $X^1_{[\infty, k_j - 1]} \cap X^1_{[k_j, \infty)} = \emptyset$ for $0 \leq j < 2J + 1$, then for any $1 \leq j \leq 2J$,

$$\emptyset \neq X^1_{(-\infty, k_j - 1]} \cap X^1_{[k_j, \infty)} = X^1_{[k_j - 1, k_j - 1]} \cap X^1_{[k_j, k_j + 1 - 1]}.$$ 

Hence using independence, we see that

$${a}_2 \leq P \left[ X^1_{[-L, -1]} \cap X^1_{[0, L - 1]} \neq \emptyset \right] \leq P[0 \notin D]^J.$$ 

(1.15)

We now turn to the control of $a_1$. We observe that

$${a}_1 \leq (2J + 1) P \left[ X^1_{(-\infty, -1]} \cap X^1_{[L, \infty)} \neq \emptyset \right] \leq (2J + 1) \sum_{i \geq 1, j \geq L} P \left[ X^1_{i+j} = 0 \right] \leq (2J + 1) \sum_{k \geq L} k P \left[ X^1_{k} = 0 \right] \leq (2J + 1) \text{const} L^{\frac{d_1 - 4}{2}},$$

(1.16)

using [11], p. 75, in the last step. Choosing a large enough $\gamma$ depending on $d_1$, $q(\cdot)$, and setting $J = \lceil \gamma \log n \rceil$, $L = \lceil \frac{\gamma}{2} \rceil$, (1.12) now follows from (1.15), (1.16). □

We will now provide an alternative representation of the law of the walk under the annealed measure $P_0$. We let $W = (\mathbb{Z}^d)^N$ stand for the space of $\mathbb{Z}^d$-valued trajectories $(\tilde{w}(k))_{k \geq 0}$ and

$$\mathcal{I}(w) = \left\{ k \geq 0, X^1_k(w) = X^1_{k+1}(w) \right\}, \quad \text{for } w \in W,$$

(1.17)

denote the non-negative idle times of $X^1$. We specify a probability kernel $K(w, d\tilde{w}d\omega)$ from $W$ to $\tilde{W} \times \Omega$ through:

$$\begin{cases}
\omega \text{ is } \mathbb{P}\text{-distributed,} \\
\tilde{w}(0) = 0, \\
\text{for any } k \geq 0, \text{ conditionally on } \omega, \tilde{w}(0), \ldots, \tilde{w}(k), \\
\tilde{w}(k+1) - \tilde{w}(k) \text{ equals } 0, \text{ when } k \geq T^1 \text{ or } k \notin \mathcal{I}(w), \\
e, \text{ with probability } \frac{\omega(X^1_k + \tilde{w}(k), e)}{q(0)}, \text{ for any} \\
e = \pm e_i, i > d_1, \text{ if } k < T^1 \text{ and } k \in \mathcal{I}(w).
\end{cases}$$

(1.18)

We can then consider the spaces

$$\Gamma_0 = W \times (\tilde{W} \times \Omega)^3 \quad \text{and} \quad \Gamma_s = W \times (\tilde{W} \times \Omega)^Z,$$

(1.19)

defined with their product $\sigma$-fields, (the subscript “0” refers to $P_0$ and the subscript “s” to stationary) and the probabilities

$$Q_0 = P \times M_0, \quad Q_s = P \times M_s,$$

(1.20)
where $M_0$ and $M_s$ stand for the respective kernels from $W$ to $(\tilde{W} \times \Omega)^N$ and $W$ to $(\tilde{W} \times \Omega)^Z$ defined by

$$M_0(w, d\gamma_0) = K(w, d\tilde{w}_0 d\omega_0) \otimes \bigotimes_{m \geq 1} K(\theta_T w, d\tilde{w}_m d\omega_m),$$

(with $\gamma_0 = (w, \gamma_0) = (w, (\tilde{w}_m, \omega_m)_{m \geq 0})$), and similarly

$$M_s(w, d\gamma_s) = \bigotimes_{m \in \mathbb{Z}} K(\theta_T w, d\tilde{w}_m d\omega_m),$$

(with $\gamma_s = (w, \gamma_s) = (w, (\tilde{w}_m, \omega_m)_{m \in \mathbb{Z}})$). We will shortly see that $(s \tau \alpha M_0, Q_0)$ is helpful in providing a representation of $X$, under $P_0$, whereas $(s \tau \alpha M_s, Q_s)$ possesses useful stationarity properties.

We now define on $(\tau \alpha M_0)$ the $\mathbb{Z}^d$-valued process $X^2_k, k \geq 0$, via

$$(1.23) \begin{cases} X^2_0 = 0, & X^2_k = \tilde{w}_0(k), \quad \text{for } 0 \leq k \leq T^1, \\ X^2_{(T^m+k),T^m+1} = X^2_T + \tilde{w}_m(k \wedge (T^{m+1} - T^m)) \quad \text{for } m \geq 1, k \geq 0, \end{cases}$$

in the notations of (1.21). In the sequel we will especially be interested in the $\mathbb{Z}^d$-valued process defined on $(\tau \alpha M_0)$:

$$Z_k = X^1_k + X^2_k, \quad k \geq 0,$$

(1.24) and by the $P_0(\cdot)$-valued process (see (0.2)):

$$\sigma_k = \omega_0(Z_k, \cdot), \quad \text{when } 0 \leq k \leq T^1,$n

$$\sigma_m(Z_k - Z_{T^m}, \cdot), \quad \text{when } T^m \leq k < T^{m+1}, m \geq 1.$$ (1.25)

The above processes will easily be compared with the processes defined on $(\tau \alpha)$:

$$Z^s_k = X^1_k + X^2_s, \quad k \in \mathbb{Z},$$

(1.26)

$$\sigma^s_m = \omega_m(Z^s_k - Z^s_{T^m}, \cdot), \quad \text{for } T^m \leq k < T^{m+1},$$

(1.27)

in the notations of (1.22), with

$$X^2_0 = 0 \quad \text{and} \quad X^2_{(T^m+k),T^m+1} = X^2_{T^m} + \tilde{w}_m(k \wedge (T^{m+1} - T^m)), \quad \text{for } m \in \mathbb{Z}, k \geq 0.$$ (1.28)

The next two propositions clarify the interest of the above objects.

**Proposition 1.2.** Under $Q_0$, $(Z_k, \sigma_k)_{k \geq 0}$ has the same law as $(X_k, \omega(X_k, \cdot))_{k \geq 0}$ under $P_0$.

**Proof.** For $\omega \in \Omega$, the $\mathbb{Z}^d$-projection of $X$, under $P_{0,\omega}$ has same law as $(X^1_k)_{k \geq 0}$ under $P$. Further for $\omega \in \Omega$ if $Y_k, k \geq 0$, is a $\mathbb{Z}^d$-valued process such that $Y_0 = 0$ and for $k \geq 0$, conditionally on $X^1, Y_0, \ldots, Y_k$, the increment $Y_{k+1} - Y_k$ is

$$\begin{cases} 0, & \text{when } k \notin \mathcal{I}(u), \\ \text{takes the value } e \text{ with probability } \frac{\omega(X^1_k + Y_k, e)}{q(0)}, & \text{for } e = \pm e_i, i > d_i, \end{cases}$$

(1.29)

when $k \in \mathcal{I}(u)$,
then
\[(X_k^1 + Y_k, \omega(X_k^1 + Y_k, \cdot))_{k \geq 0}\]
is distributed as \((X_k^1, \omega(X_k^1, \cdot))_{k \geq 0}\) under \(P_0, \omega\).
\[(1.30)\]
Letting \((\omega(x, \cdot))_{x \in \mathbb{Z}^d}\) be i.i.d. \(\mu\)-distributed (see below (0.2)), and replacing \(P_0, \omega\) with \(P_0\) the above identity of laws holds true as well. But the subsets of \(\mathbb{Z}^d\): \(X_{[0,T_1-1]}, X_{[T_1,T_2-1]}, \ldots, X_{[T_m,T_{m+1}-1]}, \ldots\) are disjoint. Hence if \((\omega_m)_{m \geq 0}\) is an i.i.d. sequence with common distribution \(\mathbb{P}\), and one replaces in (1.29), and in the first expression of (1.30) \(\omega\) with \(\omega_0\), if \(0 \leq k < T_1\) and \(\omega_m(\cdot - (X_{T_m}^1 + Y_{T_m}), \cdot), \) if \(T_m \leq k < T_{m+1}\), the identity in law is still preserved. Our claim now follows straightforwardly.

\[\square\]

To take advantage of the stationarity property on \((\Gamma, Q_\cdot)\), we introduce on \(\Gamma_{\cdot}\) the flow \((\Theta_{\cdot})_{k \in \mathbb{Z}}\) via:
\[\Theta_{\cdot}(\gamma) = (\theta_{k} w, (\tilde{w}_{n+m}, \omega_{n+m})_{m \in \mathbb{Z}})\]
on \(T_n(w) \leq k < T_{n+1}(w)\),
\[(1.31)\]
with \(\gamma\) as below (1.22). This is the natural flow extending \((\theta_{k})_{k \in \mathbb{Z}}\), if one views \((w_m, \omega_m)\) as marks of the \(\delta_{T_m}\), for \(m \in \mathbb{Z}\).

**Proposition 1.3.** –
\[Z_{n}^1 = \sum_{k=0}^{n-1} Z_{k}^1 \circ \Theta_{\cdot}, \text{ for } n \geq 1,\]
\[(1.32)\]
\[\sigma_{n}^1 = \sigma_{0}^1 \circ \Theta_{\cdot}, \text{ for } n \in \mathbb{Z},\]
\[(1.33)\]
\(\Theta_{\cdot}\) preserves \(Q_\cdot\) and in fact \((\Gamma_{\cdot}, \Theta_{\cdot}, Q_{\cdot})\) is ergodic.
\[(1.34)\]

**Proof.** – Both (1.32) and (1.33) follow by direct inspection using (1.26)–(1.28). The fact \(\Theta_{\cdot}\) preserves \(Q_\cdot\), is checked by a straightforward calculation. Let us show the ergodicity of \((\Gamma_{\cdot}, \Theta_{\cdot}, Q_{\cdot})\). The Palm measure
\[\hat{Q}_\cdot \equiv Q_\cdot(\cdot | 0 \in D) = \hat{P} \times M_{\cdot}\]
attached to the stationary point process \(N\) preserves
\[\Theta = \Theta_{\cdot}\]
\[(1.36)\]
(see Neveu [9], p. 338), and the analogue of (1.11) with \(\Theta, Q_\cdot, \hat{Q}_\cdot\) in place of \(\theta, P, \hat{P}\) and \(f\) bounded measurable holds as well. Our claim is equivalent to the ergodicity of \((\Gamma_{\cdot} \cap \{0 \in D\}, \Theta, \hat{Q}_\cdot)\). Let \(A\) be measurable subset of \(\Gamma_{\cdot} \cap \{0 \in D\}\) invariant under \(\hat{\Theta}\) and \(\varepsilon > 0\). We can find an integer \(m_{\varepsilon} \geq 1\) and a measurable subset \(A_{\varepsilon}\) depending only on \(w, (w_m, \omega_m)_{m \leq m_{\varepsilon}}\), such that:
\[E^{\hat{Q}_\cdot}[|1_{A} - 1_{A_{\varepsilon}}|] \leq \varepsilon,\]
\[(1.37)\]
Then for \(L \geq 0,\)
\[\hat{Q}_\cdot(A) = E^{\hat{Q}_\cdot}[1_{A} 1_{\Theta L}] = E^{\hat{Q}_\cdot}[1_{A_{\varepsilon}} 1_{\Theta L}] + c_{\varepsilon},\]
\[(1.38)\]
with $|c_\varepsilon| \leq 2\varepsilon$. On the other hand if $L > 2m_\varepsilon$, conditioning on the $w$ component and using the fact that the $(\tilde{w}_m, \omega_m)_{m \in \mathbb{Z}}$ are independent conditionally on $w$ (see (1.22)), the above equals

$$E^{\hat{P}}[\hat{Q}_s(A_\varepsilon | w) \hat{Q}_s(A_\varepsilon | w) \circ \hat{\theta}_L] + c_\varepsilon.$$ 

As a result

$$2\varepsilon \geq \lim_{N \to \infty} \left| \hat{Q}_s(A) - \frac{1}{N} \sum_{L=0}^{N-1} E^{\hat{P}}[\hat{Q}_s(A_\varepsilon | w) \hat{Q}_s(A_\varepsilon | w) \circ \hat{\theta}_L] \right|,$$

(1.39)

but $(W \cap \{0 \in D\}, \hat{\theta}, \hat{P})$ is ergodic as a consequence of the ergodicity of $(W, \theta, P)$ and

$$\frac{1}{N} \sum_{L=0}^{N-1} Q_s(A_\varepsilon | w) \circ \hat{\theta}_L \overset{L^1(\hat{P})}{\longrightarrow} \hat{Q}_s(A_\varepsilon).$$

We thus find with (1.37) and the above that

$$\left| \hat{Q}_s(A) - \hat{Q}_s(A)^2 \right| \leq \left| \hat{Q}_s(A) - \hat{Q}_s(A_\varepsilon) \right|^2 + 2\varepsilon \leq 4\varepsilon.$$

Letting $\varepsilon$ tend to 0, we see that $\hat{Q}_s(A) = 0$ or 1, and our claim follows. 

We will now apply the above to the derivation of a law of large numbers. In particular this will prove the existence of a (possibly vanishing) asymptotic velocity for the walk under the annealed measure $P_0$, when the single site distribution $\mu$ is concentrated on $P_q(\cdot)$, (see (0.1), (0.2), with $d_1 \geq 5, d_2 \geq 1$).

**Theorem 1.4.** — Let $\Psi$ be a bounded measurable function on $P_q(\cdot)$, then

$$P_0\text{-a.s., } \frac{1}{n} \sum_{k=0}^{n-1} \Psi(\omega(X_k, \cdot)) \overset{n \to \infty}{\longrightarrow} E^{Q_s}[\Psi(\sigma_0^s)],$$

(1.40)

and moreover in the notation of (1.1),

$$P_0\text{-a.s., } \frac{X_n}{n} \overset{d}{\longrightarrow} E^{Q_s}[d(\sigma_0^s)] = E^{Q_s}[Z_1^s].$$

(1.41)

**Proof.** — In view of Proposition 1.2, it suffices to prove similar statements with $(Z_k)_{k \geq 0}$ and $(\sigma_k)_{k \geq 0}$ in place of $(X_k)_{k \geq 0}$ and $(\omega(X_k, \cdot))_{k \geq 0}$.

In the notations of (1.19), we consider the kernel $M$ from $W$ to $(\tilde{W} \times \Omega) \times (\tilde{W} \times \Omega)^\mathbb{Z}$:

$$M(w, d\gamma) = K(w, d\tilde{w}_0 d\omega_0) \otimes \bigotimes_{m \in \mathbb{Z}} K(\theta_{\tau_0} w, d\tilde{w}_m d\omega_m),$$

(1.42)

for $\gamma = (w, \gamma) = (w, (\tilde{w}_0, \omega_0), (\tilde{w}_m, \omega_m)_{m \in \mathbb{Z}})$, and the probability $Q$ on the space $\Gamma = W \times (\tilde{W} \times \Omega) \times (\tilde{W} \times \Omega)^\mathbb{Z}$ defined as the semi-direct product $Q = P \times M$. Then the applications

$$\gamma \in \Gamma \overset{\mathbb{N}_0}{\longrightarrow} \gamma_0 = (w, (\tilde{w}_0, \omega_0), (\tilde{w}_m, \omega_m)_{m \geq 1}) \in \Gamma_0$$

$$\gamma \in \Gamma \overset{\mathbb{N}}{\longrightarrow} \gamma_s = (w, (\tilde{w}_m, \omega_m)_{m \in \mathbb{Z}}) \in \Gamma_s,$$
respectively map $Q$ onto $Q_0$ and $Q_s$. Moreover with a slight abuse of notations, we see that

$$Q\text{-a.s.}, \quad Z_{T_1+k} - Z_{T_1} = Z_{T_1+k}^r - Z_{T_1}^r, \quad \sigma_{k+T_1} = \sigma_{k+T_1}^r, \quad k \geq 0. \quad (1.43)$$

As a result we find that for $\Psi$ as in (1.40)

$$Q\text{-a.s.}, \quad \frac{1}{n} \sum_{k=0}^{n-1} \Psi(\sigma_k) - \frac{1}{n} \sum_{k=0}^{n-1} \Psi(\sigma_k^r) \to 0. \quad (1.44)$$

In view of Proposition 1.3 we can apply the ergodic theorem to the second expression in (1.44), and (1.40) follows. By (1.43), we also see that

$$Q\text{-a.s.}, \quad |Z_n - Z_n^r| \leq 2 (T_1 \wedge n), \quad (1.45)$$

and from Proposition 1.3 and the ergodic theorem we conclude that

$$P_0\text{-a.s.}, \quad \frac{X_n}{n} \to E^{Q_s}[Z_1^r]. \quad (1.46)$$

Moreover by a martingale argument (under $P_{0,\omega}$),

$$E_0[X_n] = E_0\left[\sum_{k=0}^{n-1} d(\omega(X_k, \cdot))\right]. \quad (1.47)$$

and by (1.40) we now conclude that

$$E^{Q_s}[Z_1^r] = E^{Q_s}[d(\sigma_0^r)],$$

finishing the proof of Theorem 1.4. \qed

2. Central limit theorem under the annealed measure

In the setting of the previous sections, we now present two central limit theorems for the walk under the measure $P_0$. Theorem 2.1 requires a symmetry assumption on the law of the environment, cf. (2.1) below, and holds when $d_1 \geq 7$, on the other hand Theorem 2.2 makes no symmetry assumption, but holds when $d_1 \geq 13$. We will later use Theorem 2.2 when providing in Sections 4 and 5 examples of diffusive behavior of the walk in biased environments.

For the first theorem, we assume the following “antipodal symmetry” of the single site distribution (see below (0.2))

$$\mu \text{ is invariant under } (p(e))_{|e|=1} \to (p(-e))_{|e|=1}. \quad (2.1)$$

Note that when (2.1) holds, $E_0[X_n] = 0$, for $n \geq 0$, and the limiting velocity $v$ in (1.41) necessarily vanishes. In what follows we denote by $D(\mathbb{R}_+, \mathbb{R}^d)$ the set of $\mathbb{R}^d$-valued functions on $\mathbb{R}_+$, which are right continuous with left limits, which is tacitly
endowed with the Skorohod topology and its Borel $\sigma$-algebra, (cf. Chapter 3 of Ethier and Kurtz [6]).

**Theorem 2.1** ($d_1 \geq 7$, under (2.1)). Under $P_0$, the $D(\mathbb{R}_+, \mathbb{R}^d)$-valued sequence $B_n = \frac{1}{\sqrt{n}} X_{[n]}$ converges in law to a Brownian motion with covariance matrix $A$ given in (2.14).

**Proof.** In view of Proposition 1.2 and (1.45), it suffices to show that under $Q_s$, $\frac{1}{\sqrt{n}} Z_{[n]}^i$ converges in law to a Brownian motion with covariance matrix $A$.

Define the non-decreasing sequence $k_n$, $n \geq 0$, $Q_s$-a.s. surely tending to infinity such that $T^k \leq n < T^{k+1}$, and

$$\Sigma_m = Z_{T^m}^i - Z_{T^0}^i, \quad \text{for } m \geq 0. \quad (2.3)$$

Note that $Q_s$-a.s., for any $T > 0$:

$$\sup_{t \leq T} \left| \frac{1}{\sqrt{n}} Z_{[n]}^i - \frac{1}{\sqrt{n}} \Sigma_{k_{[n]}} \right| \leq 2 \sup_{0 \leq k \leq [k_{[n]}]} \frac{(T^{k+1} - T^k)}{\sqrt{n}}. \quad (2.4)$$

From (1.12) and $d_1 \geq 7$, we see that for $\gamma < \frac{3}{2}$,

$$E^P \left[ (T^1)^\gamma \right] < \infty \quad (2.5)$$

and using (1.11) we conclude that for $\gamma < \frac{5}{2}$,

$$E^\hat{P} \left[ (T^1)^\gamma \right] = E^\hat{Q} \left[ (T^1)^\gamma \right] < \infty. \quad (2.6)$$

Using stationarity, we see that for $u > 0$,

$$P \left( \sup_{0 \leq k \leq [T_{[n]}]} \frac{(T^{k+1} - T^k)}{\sqrt{n}} > u \right) \leq \frac{(Tn + 1)}{n} \hat{P} \left( T^1 > \sqrt{n}u \right) \leq \frac{(Tn + 1)}{n} E^\hat{P} \left[ (T^1)^2 \right] \left( T^1 > \sqrt{n}u \right) \rightharpoonup 0. \quad (2.6)$$

On the other hand $\sup_{0 \leq k \leq [T_{[n]}]} \frac{(T^{k+1} - T^k)}{\sqrt{n}}$ is invariant under $\theta_{T_0}$, and by (1.11) the image of $P$ under $\theta_{T_0}$ is $T^1 \hat{P} / \int T^1 d \hat{P}$, so that the above calculation also proves that

$$\sup_{0 \leq k \leq [T_{[n]}]} \frac{(T^{k+1} - T^k)}{\sqrt{n}} \rightharpoonup 0 \quad \text{in } P \text{ (or } Q_s\text{-probability).} \quad (2.7)$$

Since $Q_s$-a.s., $k_n \leq n$ for all $n$, we see from (2.4), (2.7) that our claim will follow if we show (2.2) with $\frac{1}{\sqrt{n}} \Sigma_{k_{[n]}}$ in place of $\frac{1}{\sqrt{n}} Z_{[n]}^i$. 

---
Observe then that conditionally on \( w \), under \( \hat{Q}_s \), the variables \( Z_{T_{k+1}}^s - Z_{T_k}^s, k \geq 0 \), are independent, cf. (1.22), (1.26), (1.28), with zero mean thanks to (2.1). Further from the ergodic theorem:

\[
\hat{Q}_s\text{-a.s., } \frac{1}{n} \sum_{0 \leq k < n} (Z_{T_{k+1}}^s - Z_{T_k}^s)(Z_{T_{k+1}}^s - Z_{T_k}^s)^t \to E \hat{Q}_s \left[ (Z_{T_1}^s)(Z_{T_1}^s)^t \right] = \tilde{A}. \tag{2.8}
\]

Using the martingale central limit theorem, see Durrett [4], p. 374, or Ethier and Kurtz [6], p. 340, it follows from (2.6), (2.8) that

for \( \hat{P} \)-a.e. \( w \), conditionally on \( w \) under \( \hat{Q}_s \),

\[
\frac{1}{\sqrt{n}} \Sigma_m \cdot n \text{ converges in law to a Brownian motion with covariance matrix } \tilde{A},
\]

stands for the linear interpolation of \( \Sigma_m, m \geq 0 \).

Noting that \( \frac{1}{\sqrt{n}} \Sigma_m \) is invariant under \( \Theta_{T_0} \) and the image of \( Q_s \) under \( \Theta_{T_0} \) is \( T^1 \hat{Q}_s / \int T^1 d\hat{P} \), it follows that

under \( Q_s \), \( \frac{1}{\sqrt{n}} \Sigma_m \) converges in law to a Brownian motion with covariance matrix \( \tilde{A} \).

From the ergodic theorem, we know that

\[
\frac{T^n}{m} \to E\hat{P}[T^1] \quad \hat{Q}_s\text{-a.s.,} \tag{2.11}
\]

and by similar arguments as above the same holds true \( Q_s \)-a.s. It then follows that \( Q_s \)-a.s. \( \frac{1}{\sqrt{n}} \Sigma_{k[n]} \to 1 / \int T^1 d\hat{P} \), and with the help of Dini’s theorem:

\[
Q_s\text{-a.s., for all } T > 0, \sup_{0 \leq t \leq T} \left| \frac{k[n]}{n} - \frac{t}{E\hat{P}[T^1]} \right| = 0. \tag{2.12}
\]

From (2.10) and (2.12), we then conclude that

under \( Q_s \), \( \frac{1}{\sqrt{n}} \Sigma_{k[n]} \) converges in law to a Brownian motion

with covariance matrix

\[
A = E\hat{Q}_s \left[ (Z_{T_1}^s)(Z_{T_1}^s)^t \right] / E\hat{P}[T^1] = \tilde{A} / E\hat{P}[T^1], \tag{2.13}
\]

which finishes the proof of our claim.

We now turn to the second theorem which does not require the symmetry assumption (2.1), and covers situations with possibly non-vanishing limiting velocity \( v \), see (1.41).

**Theorem 2.2** \( (d_1 \geq 13) \). – Under \( P_0 \), the \( D(\mathbb{R}_+, \mathbb{R}^d) \)-valued sequence \( B^n = \frac{1}{\sqrt{n}}(X_{[n]} - [-n]v) \) converges in law to a Brownian motion with covariance matrix \( A \) given in (2.20).
Proof. – By Proposition 1.2 and (1.45), it suffices to prove a similar result for the sequence
\[
\frac{1}{\sqrt{n}}(Z_{\lfloor n \rfloor} - \lfloor n \rfloor v) \overset{(1.32)=(1.41)}{=} \frac{1}{\sqrt{n}} \sum_{k=0}^{\lfloor n \rfloor - 1} Y \cdot \Theta_k,
\]
with the notation
\[
Y = Z_1 - E^Q [Z_1].
\]
We now introduce on \(\Gamma_s\), see (1.19), the filtration
\[
\mathcal{G}_k = \sigma(Z_{n+1} - Z_n, n < k), \quad \text{for } k \geq 0,
\]
(\(= \sigma(Z_n, n \leq k)\), since \(Z_0 = 0\)).

The main step in proving Theorem 2.2 is provided by an adaptation of Gordin’s method:

**Lemma 2.3.** – There is a \(G \in L^2(\Gamma_s, \mathcal{G}_0, Q_s)\) such that
\[
M_n \overset{\text{def}}{=} G \circ \Theta_n - G + Z_n - n v = G \circ \Theta_n - G + \sum_{k=0}^{n-1} Y \circ \Theta_k \text{ is a } (\mathcal{G}_n)\text{-martingale}.
\]

Let us for the time being admit Lemma 2.3 and explain how we conclude the proof of Theorem 2.2. Observe that for any \(\varepsilon > 0:\)
\[
Q_s \left( \sup_{1 \leq m \leq n} |G \cdot \Theta_m| > \varepsilon \sqrt{n} \right) \leq n Q_s (|G| > \varepsilon \sqrt{n}) \leq e^{-2} E^Q [G^2, |G| > \varepsilon \sqrt{n}] \xrightarrow{n \to \infty} 0,
\]
so that it suffices to prove that \(\frac{1}{\sqrt{n}}M_{\lfloor n \rfloor}\) converges in law to conclude that \(\frac{1}{\sqrt{n}}(Z_{\lfloor n \rfloor} - \lfloor n \rfloor v)\) converges in law to the same limit. However
\[
M_n = \sum_{k=0}^{n-1} (G \circ \Theta_1 - G + Y) \circ \Theta_k
\]
is a martingale with stationary increments and from the theorem of Billingsley and Ibragimov, see Durrett [4], p. 375, it follows that
\[
\text{under } Q_s, \quad \frac{1}{\sqrt{n}}M_{\lfloor n \rfloor}\text{ converges in law to a Brownian motion with covariance matrix } A = E^Q [(G \circ \Theta_1 - G + Y)(G \cdot \Theta_1 - G + Y)'],
\]
which proves Theorem 2.2.

**Proof of Lemma 2.3.** – To simplify notations, we drop the superscript \(Q_s\) when writing expectations or conditional expectations. It follows from (1.12) that
\[
T^1 \in L^4(Q_s) \quad \text{(or } L^4(P))\].

As we now explain the claim will follow once we show that
\[ \sum_{p \geq 0} \|E[(H1 \{0 \in D\}) \circ \Theta_p | G_0]\|_2 < \infty, \] (2.22)
where we recall the notation (1.4) and
\[ H = \sum_{k=0}^{T^1-1} Y \circ \Theta_k \] (note that \(|H| \leq 2T^1\), since \(|Z^1_s| \leq 1\)). (2.23)
Indeed, if we define for \(m \geq 1\),
\[ G^m = E[H | G_0] + \sum_{1 \leq p < m} E[(H1 \{0 \in D\}) \cdot \Theta_p | G_0], \] (2.24)
then \(G^m\) converges in \(L^2(Qs)\) towards \(G \in L^2(\Gamma_s, G_0, Qs)\). Moreover, for \(m \geq 1\), we can define in the notation of (1.5),
\[ N = N(w, [1, m - 1]) + 1, \] so that for \(n \geq 0\),
\[ G \circ \Theta_n = \lim_{m \to \infty} E\left[ \left( \sum_{k=0}^{N-1} Y \circ \Theta_k \right) \circ \Theta_n | G_n \right], \] (2.25)
where the limit holds in \(L^2\) and we have used stationarity. Hence
\[ E\left[ G \circ \Theta_{n+1} + \sum_{k=0}^{n} Y \circ \Theta_k - G \circ \Theta_n - \sum_{k=0}^{n-1} Y \circ \Theta_k | G_n \right] \]
\[ = \lim_{m \to \infty} E\left[ \left( \sum_{k=0}^{T^N-1} Y \circ \Theta_k \right) \circ \Theta_{n+1} | G_{n+1} \right] + Y \circ \Theta_n \]
\[ - \left( \sum_{k=0}^{T^N-1} Y \circ \Theta_k \right) \circ \Theta_n | G_n \]
\[ = \lim_{m \to \infty} E\left[ \left( \sum_{k=0}^{T^N-1} Y \circ \Theta_k \right) \circ \Theta_{n+1} + Y \circ \Theta_n - \left( \sum_{k=0}^{T^N-1} Y \circ \Theta_k \right) \circ \Theta_n | G_n \right]. \] (2.26)
The quantity under the conditional expectation in the above expression equals
\[ 1\{n + m \in \mathcal{D}\} H \circ \Theta_{n+m} = (H1 \{0 \in D\}) \circ \Theta_{n+m} \] (2.27)
and using (2.22) and stationarity we see that the last line of (2.26) vanishes. This proves that \(M_n\), with the notation of (2.18), is a \((G_n)\)-martingale.
We are thus reduced to proving (2.22). To this end, we consider \(B \in L^2(\Gamma_s, G_0, Q_s)\) with unit \(L^2\)-norm. Then for \(p \geq 1\),
\[ E[(H1 \{0 \in D\}) \circ \Theta_p B] = \sum_{m \geq 1} E\left[ \sum_{T^m \leq k < T^{m+1}} Y \circ \Theta_k \right] B, T^m = p \] (2.28)
Note that $B$ is $\mathcal{G}_0^+$-measurable and hence a function of $w$ and $(\tilde{w}_m, \omega_m)_{m \leq 0}$, and

$$\sum_{T^m \leq k < T^{m+1}} Y \circ \Theta_k = Z^T_{T^{m+1}} - Z^T_{T^{m}} - (T^{m+1} - T^m) v$$

\[(1.26)\text{(1.28)}
X^1_{T^{m+1}} - X^1_{T^{m}} + \tilde{w}_m (T^{m+1} - T^m) - (T^{m+1} - T^m) v.
\]

Hence conditioning on $w$ in the right member of (2.28), and using the notation of (1.22), we find that for $p \geq 1$:

$$E \left[ (H1 \{0 \in \mathcal{D}\}) \circ \Theta_p B \right] = \sum_{m \geq 1} E^p \left[ (M_s H) \circ \theta_p M_s B, T^m = p \right]$$

\[= E^p \left[ (M_s H)1 \{0 \in \mathcal{D}\} \circ \theta_p M_s B \right]. \tag{2.29} \]

Then observe that we can find measurable functions $\psi$ and $\varphi$ such that

$$M_s B = \psi(T^0, (X^1_i)_{i \leq 0}), \quad (M_s H)1 \{0 \in \mathcal{D}\} = \varphi(T^1, (X^1_i)_{i \geq 0})1 \{0 \in \mathcal{D}\}. \tag{2.30}$$

To take advantage of decoupling effects, we define

$$L = \left\lfloor \frac{p}{3} \right\rfloor, \tag{2.31}$$

and introduce two copies $(X^-_n)$ and $(X^+_n)$ of $(X^1_n)$, such that $X^-_n$ coincides with $X^1_n$ for $n \leq L$ and then “evolves” independently, whereas $X^+_n$ coincides with $X^1_{n+1}$ for $n \geq -L$, and for $n < -L$, “evolves” independently. We then define

$$\begin{align*}
U &= M_s B, \quad U^- = \psi(T^-, (X^-_i)_{i \leq 0}), \\
V &= ((M_s H)1 \{0 \in \mathcal{D}\}) \circ \theta_p = \psi(T^1 \circ \theta_p, (X^1_{i+p} - X^1_p)_{i \geq 0})1 \{p \in \mathcal{D}\}, \\
V^+ &= \varphi(T^+, (X^+_i)_{i \geq 0})1 \{0 \in \mathcal{D}^+\},
\end{align*} \tag{2.32}$$

where $T^-$ and $T^+$ are respectively defined like $T^0$ and $T^1$ relatively to $(X^-)$ and $(X^+)$ and $\mathcal{D}^+$ is defined analogously to $\mathcal{D}$ with $(X^+)$ in place of $X^-$. We of course tacitly abuse the notations since the above objects are defined on an extension of the space $(\mathcal{W}, \mathcal{W}, P)$. Note that

$$U \overset{\text{law}}{=} U^-, \quad V \overset{\text{law}}{=} V^+. \tag{2.33}$$

We now find that for $p \geq 1$:

\begin{align*}
E \left[ (H1 \{0 \in \mathcal{D}\}) \circ \Theta_p B \right] &\overset{(2.29)}{=} E^p [V U] \\
&= E^p [V^+ U^-] + E^p [V^+ (U - U^-)] + E^p [(V - V^+) U]. \tag{2.34}
\end{align*}

Note also that:

$$E^p [V] = E^p [V^+] = E \left[ H1 \{0 \in \mathcal{D}\} \right] = E[Y] E^p [T_1] = 0, \tag{2.35}$$

using the analogue of (1.11) for $Q_s$, $\hat{Q}_s$, and (2.16) in the third equality. Note that $V^+$ and $U^-$ are independent. Hence the first term in the last member of (2.34) vanishes. Keeping
in mind that $B$ has unit $L^2$-norm we find
\[
|E[(H1|0 \in \mathcal{D}) \circ \Theta_p B]| \leq \|V^+(U - U^-)\|_1 + \|V - V^+\|_2. \tag{2.36}
\]
In view of (2.32) and the inequality $|H| \leq 2T^1$, we find
\[
|V| \leq 2|T^1 \circ \theta_p|, \quad |V^+| \leq 2|T^+|,
\]
\[
|V - V^+| \leq 2(1\{T^+ \neq T^1 \circ \theta_p\} + |1\{p \in \mathcal{D}\} - 1\{0 \in \mathcal{D}^+\}|)(T^+ + T^1 \circ \theta_p). \tag{2.37}
\]
Using Cauchy–Schwarz’s inequality and stationarity, we find
\[
\|V - V^+\|_2 \leq 4\|T^1\|_2 (p[T^+ \neq T^1 \circ \theta_p])^{1/2} + 2P[\{p \in \mathcal{D}\}\{0 \in \mathcal{D}^+\}]^{1/2}. \tag{2.38}
\]
Since $X_n^+ + X_{n+p}^1 - X_p^1$ coincide for $n \geq -L$, we see that:
\[
\{T^1 \circ \theta_p \neq T^+\}
\subseteq \{\{X^1_{(\infty,-L] } \cap X_{\infty}^1 \neq \emptyset\} \cup \{X^1 \circ \theta_p\}_{(-\infty,-L]} \cap \{X^1 \circ \theta_p\}_{(0,\infty]} \neq \emptyset\}, \tag{2.39}
\]
and by a similar argument $\{p \in \mathcal{D}\}\{0 \in \mathcal{D}^+\}$ is included in the right-hand side of (2.39). As a result we obtain:
\[
\|V - V^+\|_2 \leq 24\|T^1\|_2 P[\{X^1_{(\infty,-L] } \cap X_{\infty}^1 \neq \emptyset\}]^{1/2}. \tag{2.40}
\]
By analogous arguments we also have
\[
|U - U^-| \leq (|U| + |U^-|)\{T^0 \neq T^-\}
\leq (|U| + |U^-|)\{1\{X_{(\infty,0]}^1 \cap X_{L,\infty}^1 \neq \emptyset\} + 1\{X_{(-\infty,0]} \cap X_{L,\infty} \neq \emptyset\}). \tag{2.41}
\]
Using Hölder’s inequality and $\|U\|_2 = \|U^-\|_2 \leq 1$, we find
\[
\|V^+(U - U^-)\|_1 \leq 4\|T^1\|_2 P[\{X_{(\infty,0]} \cap X_{L,\infty}^1 \neq \emptyset\}]^{1/4}. \tag{2.42}
\]
Collecting (2.36), (2.40), (2.42), and using the fact that $(X_n^1)$ and $(X_{-L}^1)$ have same law (see (1.2)), we find
\[
\|E[(H1|0 \in \mathcal{D}) \circ \Theta_p | G_0]\|_2 \leq 28\|T^1\|_4 P[\{X_{(\infty,-L)}^1 \cap X_{(0,\infty]}^1 \neq \emptyset\}]^{1/4}. \tag{2.43}
\]
By the calculation in (1.16) we know that the rightmost expression is bounded by $\text{const } p^{-(d_1-4)/8}$, (recall (2.31)), and hence summable in $p$ since $d_1 \geq 13$. This concludes the proof of (2.22) and consequently of Lemma 2.3. \(\square\)

3. Central limit theorem under the quenched measure

In this section we will explain how the central limit theorems of the previous section can be strengthened into statements under the quenched measure $P_{0,\omega}$, for $\mathbb{P}$-a.e. $\omega$. 


THEOREM 3.1. – Assume \( d_1 \geq 7 \) and (2.1) or \( d_1 \geq 13 \). Then for \( \mathbb{P}\)-a.e. \( \omega \), under \( P_{0,\omega} \), the \( D(\mathbb{R}_+, \mathbb{R}^d) \)-valued \( B^\omega \) converges in law to a Brownian motion with covariance \( A \) given in Theorems 2.1 and 2.2 respectively.

Proof. – The claim will follow from a variance calculation. It is convenient to introduce the space \( C(\mathbb{R}^d) \) of continuous \( \mathbb{R}^d \)-valued functions on \( \mathbb{R}_+ \), and the \( C(\mathbb{R}_+, \mathbb{R}^d) \)-valued variable

\[
\beta^\omega_n = \text{the polygonal interpolation of } \frac{k}{n} \rightarrow B^\omega_k, \quad k \geq 0. \tag{3.1}
\]

It will also be useful to consider the analogously defined space \( C([0, T], \mathbb{R}^d) \), of continuous \( \mathbb{R}^d \)-valued functions on \( [0, T] \), for \( T > 0 \), which we endow with the distance

\[
d_T(v, v') = \sup_{s \leq T} |v(s) - v'(s)| \land 1. \tag{3.2}
\]

From Lemma 4.1 of [1], the claim will follow once we show that for all \( T > 0 \), for all bounded Lipschitz functions \( F \) on \( C([0, T], \mathbb{R}^d) \) and \( b \in (1, 2) \):

\[
\sum_m \operatorname{var}_P(E_{0,\omega}[F(\beta^\omega_m)]) < \infty \tag{3.3}
\]

(with a slight abuse of notations).

Before proving (3.3) we still need to introduce some further notations. Given \( \omega \in \Omega \), we consider two independent copies \( (X_k)_{k \geq 0} \) and \( (\tilde{X}_k)_{k \geq 0} \) evolving according to \( P_{0,\omega} \). The respective \( \mathbb{Z}^d \)-projections \( (X^1_k)_{k \geq 0} \) and \( (\tilde{X}^1_k)_{k \geq 0} \) are then independent and with distribution given in (1.2). We then denote by \( \mathcal{C} \) the set of one-sided cut-times of \( X^1 \):

\[
\mathcal{C} = \{ k \geq 1, \ X^1_{[0,k-1]} \cap X^1_{[k,\infty)} = \emptyset \}, \tag{3.4}
\]

with an analogously defined \( \tilde{\mathcal{C}} \) attached to \( \tilde{X}^1 \). We then pick:

\[
b \in (1, 2), \quad 0 < \mu < \nu < \frac{1}{2}, \tag{3.5}
\]

and for \( m \geq 1 \), we define \( n = [b^m] \),

\[
\tau_m = \inf \{ \mathcal{C} \cap [n^\mu, \infty) \} < \infty, \quad P_{0,\omega}\text{-a.s. (cf. Lemma 1.1),} \tag{3.6}
\]

as well as the corresponding variable \( \tilde{\tau}_m \) attached to \( \tilde{X}^1 \). We will also need the event:

\[
\mathcal{A}_m = \{ \tau_m \vee \tilde{\tau}_m \leq n^\nu \text{ and } X^1_{[\tau_m, \infty)} \cap \tilde{X}^1_{[\tilde{\tau}_m, \infty)} = \emptyset \}. \tag{3.7}
\]

We now prove (3.3). Without loss of generality, we assume that \( |F| \leq 1 \) and the Lipschitz constant of \( F \) is smaller than 1. Then for \( m \geq 1 \):

\[
\operatorname{var}_P(E_{0,\omega}[F(\beta^\omega_m)]) = \mathbb{E}[E_{0,\omega} \otimes E_{0,\omega}[F(\beta^\omega_n)F(\tilde{\beta}^\omega_n)] - E_{0} \otimes E_{0}[F(\beta^\omega_n)F(\tilde{\beta}^\omega_n)]
\]

\[
= \mathbb{E}[E_{0,\omega} \otimes E_{0,\omega}[F(\beta^\omega_n)F(\tilde{\beta}^\omega_n), \mathcal{A}_m]] + \mathbb{E}[E_{0,\omega} \otimes E_{0,\omega}[F(\beta^\omega_n)F(\tilde{\beta}^\omega_n), \mathcal{\tilde{A}}_m]] - E_{0} \otimes E_{0}[F(\beta^\omega_n)F(\tilde{\beta}^\omega_n), \mathcal{A}_m] + d_m, \tag{3.8}
\]
and with a slight abuse of notations

\[ |d_m| \leq 2(P \times P)(A_m^c). \tag{3.9} \]

Moreover observe that \( P_0 \)-a.s.

\[ \sup_{s \geq 0} |(\beta^n_{s+m/n} - \beta^n_{m/n}) - \beta^a_s| \leq \frac{2}{\sqrt{n}}(\tau_m + 1), \quad \text{and} \]

\[ \beta^n_{s+m/n} - \beta^n_{m/n} = \text{the polygonal interpolation of } k \rightarrow \frac{1}{\sqrt{n}}(X_k + \tau_m - X_{\tau_m} - kv). \tag{3.11} \]

From the Lipschitz property of \( F \) and (3.7) we see that the first two terms of the last member of (3.8) equal

\[ \mathbb{E}[E_{0,0} \otimes E_{0,0} [F(\beta^n_{s+m/n} - \beta^n_{m/n}) F(\tilde{\beta}^n_{s+m/n} - \tilde{\beta}^n_{m/n}), A_m]] \]

\[ - \mathbb{E}_{0,0} [F(\beta^n_{s+m/n} - \beta^n_{m/n}) F(\tilde{\beta}^n_{s+m/n} - \tilde{\beta}^n_{m/n}), A_m] + e_m, \quad \text{with} \]

\[ |e_m| \leq \frac{8}{\sqrt{n}}(n^\nu + 1). \tag{3.13} \]

Keeping in mind the definition of \( A_m \) in (3.7), we see by conditioning on \( X_1 \) and \( \tilde{X}_1 \) that the difference of the first two terms of (3.13) vanishes. Since clearly \( \sum_m |e_m| < \infty \), (recall \( n = [b^m] \)), we only need to observe that

\[ \sum_m (P \times P)(A_m^c) < \infty. \tag{3.14} \]

By a similar calculation as in (1.16), we see that

\[ P \times P[X_{[n^\nu, \infty]} \cap \tilde{X}_{[n^\nu, \infty]} \neq \emptyset] \leq \text{const } n^{-\mu \frac{(d-4)}{2}}, \tag{3.15} \]

moreover \( \tau_m - n^\nu \) is stochastically dominated by \( T_1 \) (under the \( P \)-measure) so that from (1.12), for large \( m \):

\[ P[\tau_m > n^\nu] \leq \text{const } (\log n^\nu)^{1 + \frac{d-4}{2}} n^{-\frac{(d-4)}{2}} \leq e^{-\text{const } m}. \tag{3.16} \]

Combining (3.15) and (3.16) we deduce (3.14). \( \square \)

4. Diffusive behavior in a slightly biased environment

As explained in the introduction, it was shown in [14], that when the single-site distribution is concentrated on \( \varepsilon \)-perturbations of the \( d \)-dimensional simple random walk and \( \mathbb{E}[d(0, \omega)] \) has size bigger than \( \varepsilon^{5/2-\eta} \), when \( d = 3, \varepsilon^{3-\eta} \), where \( d \geq 4 \), then for small \( \varepsilon \), depending on \( d \) and \( \eta \in (0, 1) \), the walk has non-vanishing limiting velocity (in fact much more is known, see [14]). In this section we provide examples of \( \varepsilon \)-perturbations of the simple random walk for which \( \mathbb{E}[d(0, \omega)] \neq 0 \), but the ballistic
behavior is lost, when \( d \geq 7 \), and the diffusive behavior is even demonstrated when \( d \geq 15 \). We keep the notations of the previous sections, and specialize here \( \kappa \) to \( \frac{1}{d} \), and \( q(\cdot) \) in (0.1) to
\[
q(e) = \begin{cases} 
\frac{d_2}{d}, & \text{if } e = 0, \\
\frac{1}{2d}, & \text{if } e = \pm e_i, \quad 1 \leq i \leq d_1.
\end{cases}
\]

(4.1)

Recall the definition of \( S_\varepsilon \) in (0.7). Note that when \( p(\cdot) \in S_\varepsilon \), \( p(e) \geq \kappa \), for all \( e \), and \( \kappa \) is a global ellipticity constant. The main object of this section is the following

**Theorem 4.1.** Assume \( d \geq 7 \), then for all \( \varepsilon \in (0, 1) \), we can find \( \mu \) concentrated on \( S_\varepsilon \), such that
\[
\mathbb{E}[d(0, \omega)] \neq 0, \quad \text{but} \quad P_0\text{-a.s., } \frac{X_n}{n} \to 0, \quad \text{as } n \to \infty.
\]

(4.2)

In addition when \( d \geq 15 \), we can make sure that for \( \mathbb{P}\text{-a.e.} \omega \),

under \( P_{0, \omega} \frac{1}{\sqrt{n}} X_{\lfloor n \rfloor} \) converges in law towards a Brownian motion with covariance matrix \( A \) (independent of \( \omega \)).

(4.3)

\[
\mathcal{P}_{q(\cdot)}^s = \left\{ p(\cdot) \in \mathcal{P}_{q(\cdot)}, \text{ such that } p(e) = p(-e) \text{ for all } e \right\},
\]

(4.5)

and define \( \Omega_0 = \left( \mathcal{P}_{q(\cdot)}^s \cap S_\varepsilon \right)^{2d} \). We will use the following

**Lemma 4.2.** Suppose \( \varphi \) is a measurable function on \( \mathcal{P}_{q(\cdot)}^s \cap S_\varepsilon \) with values in \([-1, 1]\), and \( \mu_0 \) a probability on \( \mathcal{P}_{q(\cdot)}^s \cap S_\varepsilon \) such that:
\[
\int \varphi(p) \, d\mu_0(p) = 0, \quad \text{and} \quad \mathbb{E}^{Q_0^\varepsilon}[\varphi(\sigma^\varepsilon_1)] \neq 0,
\]

(4.6)

(4.7)

where \( Q_0^\varepsilon \) denotes the probability constructed in (1.20) when the single site distribution is \( \mu_0 \). Then one can find a \( \mu \) concentrated on \( \mathcal{P}_{q(\cdot)}^s \cap S_\varepsilon \), for which \( \int d(p) \, d\mu(p) \neq 0 \), but \( v = 0 \).

**Proof.** We will look for environments of the form
\[
\omega_{p, \lambda}(x, e) = \omega_0(x, e) + \rho(\varphi(\omega_0(x, \cdot)) + \lambda) e_d \cdot e, \quad x \in \mathbb{Z}^d, \quad |e| = 1,
\]

(4.8)
with \( \rho \in [0, \frac{\epsilon}{16d}] \), \( \lambda \in [-1, 1] \), two parameters and \( \omega_0 \) distributed according to \( \mathbb{P}_0 = \mu_0^{\otimes \mathbb{Z}^d} \). The distribution \( \mu \) will correspond to the single site distribution \( \mu_{\rho, \lambda} \) of the above \( \omega_{\rho, \lambda} \) for small \( \rho \) and an appropriate choice of \( \lambda \). Note that \( \mu_{\rho, \lambda} \) is automatically concentrated on \( \mathcal{P}_q(\cdot) \cap \mathcal{S}_\epsilon \).

For \( \rho, \lambda \) as above, we consider the kernel \( K_{\rho, \lambda} \) from \( W \) to \( \tilde{W} \times \Omega_0 \), defined as in (1.18) with the difference that \( \omega \) is replaced by \( \omega_{\rho, \lambda} \), and denote by \( v_{\rho, \lambda} \) the asymptotic velocity corresponding to the single site distribution \( \mu_{\rho, \lambda} \), see (1.41). We now find that

\[
v_{\rho, \lambda} = \frac{E^{\hat{P} \times K_{\rho, \lambda}} \left[ \sum_{k=0}^{T-1} d(\omega_{\rho, \lambda}(X^1_k + \tilde{w}(k), \cdot)) \right]}{E^{\hat{P}}[T^1]},
\]

(4.8)

From the above formula one deduces that

\[(\rho, \lambda) \in \left[0, \frac{\epsilon}{16d}\right] \times [-1, 1] \rightarrow v_{\rho, \lambda} \text{ is a continuous function}. \quad (4.10)\]

Indeed given \((\rho_0, \lambda_0)\) and \((\rho_1, \lambda_1)\), one can couple the two kernels \( K_{\rho_0, \lambda_0} \) and \( K_{\rho_1, \lambda_1} \) so that when both walks are at time \( k < T^1 \) in the same location \( x \), they simultaneously jump to \( x + e \) with probability \( \omega_{\rho_0, \lambda_0}(x, e) \wedge \omega_{\rho_1, \lambda_1}(x, e) \). The asserted continuity follows then from dominated convergence. Note also by direct inspection of the last line of (4.9) that

\[
v_{\rho, 1} \cdot e_d \geq 0 \quad \text{and} \quad v_{\rho, -1} \cdot e_d \leq 0, \quad \text{for} \quad 0 < \rho \leq \frac{\epsilon}{16d}. \quad (4.11)\]

We can hence define for \( 0 < \rho \leq \frac{\epsilon}{16d} \):

\[
\lambda_{\rho} \overset{\text{def}}{=} \text{the largest zero of the continuous function} \quad \lambda \rightarrow v_{\rho, \lambda}. \quad (4.12)\]

We see that for \( 0 < \rho \leq \frac{\epsilon}{16d} \):

\[
\begin{cases}
v_{\rho, \lambda_{\rho}} = 0, \\
\lambda_{\rho} = -\frac{E^{\hat{P} \times K_{\rho, \lambda}} \left[ \sum_{k=0}^{T-1} \varphi(\omega_0(X^1_k + \tilde{w}(k), \cdot)) \right]}{E^{\hat{P}}[T^1]}, \quad (4.13) \\
\int d(p) d\mu_{\rho, \lambda_{\rho}} = 2\rho \left( \int \varphi(p) d\mu_0(p) + \lambda_{\rho} \right) e_d \overset{(4.6)}{=} 2\rho \lambda_{\rho} e_d. 
\end{cases}
\]

On the other hand a similar coupling argument as above shows that

\[
\lim_{\rho \to 0} \frac{E^{\hat{P} \times K_{\rho, \lambda_{\rho}} \left[ \sum_{k=0}^{T-1} \varphi(\omega_0(X^1_k + \tilde{w}(k), \cdot)) \right]}}{E^{\hat{P}}[T^1]} = \frac{E^{\hat{P}}[\varphi(\sigma_0^1)]}{E^{\hat{P}}[\varphi(\sigma_0^1)]}. \quad (4.14)
\]
As a result, we obtain that
\[
\lim_{\rho \to 0} \lambda_\rho = -E(Q_s|\sigma_\rho^s) \neq 0,
\]
so that for small \( \rho, \mu_\rho, \lambda_\rho \) satisfies the claims of Lemma 4.2. \( \square \)

Remark 4.3. – (1) With minor modifications one obtains a similar statement for an \( \mathbb{R}^{d_2} \)-valued \( \varphi \), with \( |\varphi| \leq 1 \), and analogous assumptions as in (4.6), (4.7). One now chooses for \( 0 \leq \rho \leq \epsilon \) and \( \lambda \) in the closed unit ball of \( \mathbb{R}^{d_2} \),
\[
\omega_\rho, \lambda(x, e) = \omega_0(x, e) + \rho (\varphi(\omega_0(x, \cdot)) + \lambda) \cdot e,
\]
in place of (4.8), and uses Brouwer’s fixed point theorem, cf. Dugundji [3], p. 341, to find \( \lambda_\rho \) for \( 0 < \rho \leq \epsilon \), satisfying the second equality of (4.13). This remark may be helpful if one wishes that the distribution \( \mu \) of Theorem 4.1 accommodates a genuinely vector-valued local drift.

(2) Analogously with a slight change in the proof of Lemma 4.2, under the same assumptions, setting \( \lambda = 0 \) in (4.8), one constructs \( \mu \) concentrated on \( P_q(\cdot) \cap S_\epsilon \) such that
\[
\int d(p) d\mu(p) = 0, \quad \text{but} \quad v \neq 0.
\]
Thus Theorem 4.1 holds as well when (4.2), (4.3), (4.4) are respectively replaced by:
\[
E[\mathbb{Q}_s|\sigma_0^s] = 0, \quad \text{P}_0 \text{-a.s.} \quad \frac{X_n}{n} \to v \neq 0, \quad (v \text{ deterministic}),
\]
under \( P_{0,\omega} \), \( \frac{1}{\sqrt{n}}(X_{[n]} - \lfloor n \rfloor v) \) converges in law towards a Brownian motion with covariance matrix \( A \) (independent of \( \omega \)). \( \quad \text{(4.4)} \)

We now proceed with the proof of Theorem 4.1. We are reduced to checking the assumptions of Lemma 4.2. To this end we will use the general

**Lemma 4.4 (under the assumptions of Section 1).** – For \( \Psi \) bounded measurable on \( \mathcal{P}_{Q_s} \)
\[
E[\Psi(\sigma_0)] = \lim_{n \to \infty} \frac{1}{n} \sum_{x \in \mathbb{Z}^d} E\left[ e_n(x, 0, \omega) \frac{\Psi(\omega(0, \cdot))}{\sum_{|e|=1} \omega(0, e) P_{e,\omega}[H_0 = \infty]} \right],
\]
with
\[
e_n(x, y, \omega) = E_{x,\omega}[u^H, H_y < \infty], \quad \text{for} \ x, y \in \mathbb{Z}^d, \ \omega \in \Omega, \ n \geq 1, \ \text{where}
\]
\[
u = 1 - \frac{1}{n} \text{ and } H_z = \inf\{k \geq 0, \ X_k = z\}, \ \text{for} \ z \in \mathbb{Z}^d.
\]

**Proof.** – We write \( S_m = \sum_{k=0}^{m} \omega(X_k, \cdot) \), for \( m \geq 0 \), and \( S_{-1} = 0 \), so that
\[
\sum_{m=0}^{\infty} u^m \omega(X_m, \cdot) = \sum_{m=0}^{\infty} u^m (S_m - S_{m-1}) = \frac{1}{n} \sum_{m=0}^{\infty} u^m S_m.
\]
Noting that \( \frac{1}{n^2} \sum_{m=0}^{\infty} m u^m = 1 - \frac{1}{n^2} \), it follows from (1.40) that:

\[
\lim_{n \to \infty} \frac{1}{n} E_0 \left[ \sum_{m=0}^{\infty} u^m \Psi(\omega(X_m, \cdot)) \right] = E_0^Q \left[ \Psi(\sigma^e_0) \right].
\] (4.18)

On the other hand for \( \omega \in \Omega \), setting

\[
g_n(x, y, \omega) = E_{x, \omega} \left[ \sum_{m \geq 0} u^m 1 \{ X_m = y \} \right],
\] (4.19)

we find

\[
E_{0, \omega} \left[ \sum_{m=0}^{\infty} u^m \Psi(\omega(X_m, \cdot)) \right] = \sum_{x \in \mathbb{Z}^d} e_n(0, x, \omega) \frac{\Psi(\omega(x, \cdot))}{1 - E_{x, \omega}[u^{H_0}]},
\] (4.20)

by a classical Markov chain calculation, provided

\[
\tilde{H}_z = \inf\{ k \geq 1, X_k = z \} \quad \text{for } z \in \mathbb{Z}^d.
\] (4.21)

Since the \( \mathbb{Z}^{d_1} \)-projection of \( X \), under \( P_{0, \omega} \), is distributed as \( X_1 \) under \( P \), we have:

\[
1 - E_{0, \omega}[u^{\tilde{H}_z}] \geq 1 - P_{0, \omega}[\tilde{H}_z < \infty] \geq P[X^1_k \neq 0, \text{ for all } k \geq 1] > 0.
\] (4.22)

Moreover for any \( |e| = 1 \),

\[
\lim_{n \to \infty} \sup_{\omega} E_{e, \omega}[u^{H_0}] - P_{e, \omega}[H_0 < \infty] = 0,
\] (4.23)

since for \( M > 0 \),

\[
0 \leq P_{e, \omega}[H_0 < \infty] - E_{e, \omega}[u^{H_0}] = E_{e, \omega}[(1 - u^{H_0}), H_0 < \infty] \\
\leq 1 - u^M + P_{e, \omega}[M < H_0 < \infty] \leq 1 - u^M + P[X^1_n = 0, \text{ for some } n \geq M],
\]

from which (4.23) follows by letting \( n \) and then \( M \) tend to infinity. From (4.20) we see by choosing \( \Psi = 1 \), that for \( \omega \in \Omega \),

\[
\sum_{x \in \mathbb{Z}^d} \frac{1}{n} e_n(0, x, \omega) \leq 1.
\] (4.24)

Integrating over the environment in (4.20) and using translation invariance, as well as (4.22), (4.23), we obtain:

\[
\lim_{n \to \infty} \frac{1}{n} E_0 \left[ \sum_{m=0}^{\infty} u^m \Psi(\omega(X_m, \cdot)) \right] = \lim_{n \to \infty} \frac{1}{n} \sum_{x \in \mathbb{Z}^d} \mathbb{E} \left[ e_n(x, 0, \omega) \frac{\Psi(\omega(0, \cdot))}{1 - E_{0, \omega}[u^{H_0}]} \right].
\]
\[
\lim_{n \to \infty} \sum_{x \in \mathbb{Z}^d} \frac{1}{n} \mathbb{E} \left[ e_n(x, 0, \omega) \frac{\Psi(\omega(0, \cdot))}{P_{0,\omega}[H_0 = \infty]} \right],
\]
which together with (4.18), finishes the proof of (4.16). \(\Box\)

The distribution \(\mu_0\) of Lemma 4.2, that we now construct, will be concentrated on small perturbations of

\[
p_\nu(e) = \begin{cases} 
\frac{1}{d}, & \text{for } e = \pm e_i, \ 1 \leq i \leq d - 2, \\
\frac{\nu}{2d}, & \text{for } e = \pm e_{d-1}, \\
\frac{2 - \nu}{2d}, & \text{for } e = \pm e_d.
\end{cases}
\]

Note that \(p_\nu(\cdot) \in \mathcal{P}^s \cap \mathcal{S}_\varepsilon\). We denote by \(P_{\nu, x}\), for \(x \in \mathbb{Z}^d\), the canonical law of the random walk with jump distribution \(p_\nu(\cdot)\), starting from \(x\). Let us admit for the time being the fact that for small \(\varepsilon\),

\[
\Delta(\varepsilon) \overset{\text{def}}{=} P_{\nu, e_d}[H_0 < \infty] - P_{\nu, e_{d-1}}[H_0 < \infty] > 0,
\]

and explain how we complete the construction of \(\mu_0\) and \(\varphi\) of Lemma 4.2. We choose \(\mu_0\) concentrated on \(\mathcal{P}^s \cap \mathcal{S}_\varepsilon\) such that

\[
\mu_0\text{-a.s., } \quad p(e) = p_\nu(e), \quad \text{for } e = \pm e_i, \ 1 \leq i \leq d - 2, \quad \text{and} \quad \tilde{\delta} \overset{\text{def}}{=} p(e_d) - p_\nu(e_d) = -(p(e_{d-1}) - p_\nu(e_{d-1})) \quad \text{is such that}
\]

\[
0 < \|\tilde{\delta}\|_\infty \leq \frac{\varepsilon}{64d}, \quad \int \tilde{\delta} \, d\mu_0 = 0, \quad \int \tilde{\delta}^2 \, d\mu_0 \geq \frac{\|\tilde{\delta}\|_\infty^2}{2}.
\]

Such a choice is of course possible. We then define

\[
\varphi(p) = \tilde{\delta},
\]

so that \(|\varphi| \leq 1\), and (4.6) is satisfied. Writing \(\tilde{\delta}(x)\) for \(\omega(x, e_d) - p_\nu(e_d), \ x \in \mathbb{Z}^d\), we deduce from (4.16) that

\[
\mathbb{E}^{\tilde{\delta}_0} \left[ \varphi(\sigma_0^0) \right] = \lim_{n \to \infty} \frac{1}{n} \sum_x \mathbb{E} \left[ e_n(x, 0, \omega) \frac{\tilde{\delta}(0)}{P_{0,\omega}[H_0 = \infty]} \right],
\]

where \(\mathbb{E}\) stands for the \(\mu_0^{\mathbb{Z}^d}\)-expectation. Note that

\[
P_{0,\omega}[H_0 = \infty] = P_{0,\omega}[\hat{H}_0 = \infty] \left( 1 + \sum_{|e| = 1} (\omega(0, e) - p_\nu(e)) \frac{P_{e,\omega}[H_0 = \infty]}{P_{0,\omega}[H_0 = \infty]} \right),
\]

where \(\overline{P}_{0,\omega}\) denotes the probability corresponding to the environment \(\overline{\omega}\), which coincides with \(\omega\) outside 0 and such that \(\overline{\omega}(0, \cdot) = p_\nu(\cdot)\). Note that the sum inside the parenthesis
in the above expression is a.s. bounded by $\frac{4\tilde{\delta}}{\kappa}$ (see also the remark below (4.1) about $\kappa$). Using the inequality $|\frac{1}{1+\gamma} - 1 + \gamma| \leq 2\gamma^2$, for $|\gamma| \leq \frac{1}{2}$, we see that for $x \in \mathbb{Z}^d$,

$$
\mathbb{E} \left[ e_n(x, 0, \omega) \frac{\tilde{\delta}(0)}{P_{0,\omega}[H_0 = \infty]} \right] = \mathbb{E} \left[ e_n(x, 0, \omega) \frac{\tilde{\delta}(0)}{P_{0,\omega}[H_0 = \infty]} \right]
$$

Using independence we see the first term in the right-hand side of (4.32) vanishes, and

$$
\mathbb{E} \left[ e_n(x, 0, \omega) \frac{\tilde{\delta}(0)}{P_{0,\omega}[H_0 = \infty]} \right] = \mathbb{E} \left[ e_n(x, 0, \omega) \frac{\tilde{\delta}(0)}{P_{0,\omega}[H_0 = \infty]} \right] - \mathbb{E} \left[ \frac{e_n(x, 0, \omega)}{P_{0,\omega}[H_0 = \infty]} \tilde{\delta}(0) \sum_{|e|=1} (\omega(0, e) - p_{e,\omega}) P_{e,\omega}[H_0 = \infty] \right] + \mathbb{E} \left[ \frac{e_n(x, 0, \omega)}{P_{0,\omega}[H_0 = \infty]} B(x, \omega) \right],
$$

and for $|B(x, \omega)| \leq \frac{32\|\tilde{\delta}\|_\infty^3}{\kappa^2}$. (4.32)

Using independence we see the first term in the right-hand side of (4.32) vanishes, and

$$
\frac{1}{n} \sum_{x \in \mathbb{Z}^d} \mathbb{E} \left[ e_n(x, 0, \omega) \frac{\tilde{\delta}(0)}{P_{0,\omega}[H_0 = \infty]} \right]
$$

Using independence we see the first term in the right-hand side of (4.32) vanishes, and

$$
\frac{1}{n} \sum_{x \in \mathbb{Z}^d} \mathbb{E} \left[ e_n(x, 0, \omega) \frac{\tilde{\delta}(0)}{P_{0,\omega}[H_0 = \infty]} \right] = -\frac{2}{n} \sum_{x \in \mathbb{Z}^d} \mathbb{E} \left[ \frac{e_n(x, 0, \omega)}{P_{0,\omega}[H_0 = \infty]} \right] \mathbb{E} \left[ \tilde{\delta}^2 \right]
$$

With $|C| \leq \frac{32\|\tilde{\delta}\|_\infty^3}{\kappa^2}$, we have

$$
\frac{1}{n} \sum_{x \in \mathbb{Z}^d} \mathbb{E} \left[ e_n(x, 0, \omega) \frac{\tilde{\delta}(0)}{P_{0,\omega}[H_0 = \infty]} \right] = \frac{1}{n} \sum_{x \in \mathbb{Z}^d} \mathbb{E} \left[ e_n(x, 0, \omega) \right] \frac{\tilde{\delta}(0)}{P_{0,\omega}[H_0 = \infty]}
$$

Note that by choosing $\|\tilde{\delta}\|_\infty$ sufficiently small, we can make sure that $\mu_{0}^{\otimes 2^d}$-a.s.

$$
P_{e_{d-1},\omega}[H_0 < \infty] - P_{e_{d-1},\omega}[H_0 < \infty] \geq \frac{1}{2} \Delta(\epsilon), \quad \text{cf. (4.27), (4.34)}
$$

so that using (4.29) as well, the first term in the left member of (4.33) is bigger than:

$$
\frac{1}{2} \Delta(\epsilon) \|\tilde{\delta}\|_\infty^2 \frac{1}{n} \sum_{x \in \mathbb{Z}^d} \mathbb{E} \left[ e_n(x, 0, \omega) \right] \frac{\tilde{\delta}(0)}{P_{0,\omega}[H_0 = \infty]}
$$

Observe that

$$
\kappa \mathbb{P}_{0,\omega}[\tilde{H}_0 = \infty] \leq P_{0,\omega}[\tilde{H}_0 = \infty] \leq \frac{1}{\kappa} \mathbb{P}_{0,\omega}[\tilde{H}_0 = \infty]
$$

and

$$
\lim \frac{1}{n} \sum_{x \in \mathbb{Z}^d} \mathbb{E} \left[ e_n(x, 0, \omega) \right] \frac{\tilde{\delta}(0)}{P_{0,\omega}[H_0 = \infty]}
$$

As a result we see that

$$
E^{\otimes 0} [\varphi (s^T_0)] \geq \frac{\kappa}{2} \Delta(\epsilon) \|\tilde{\delta}\|_\infty^2 - \frac{32\|\tilde{\delta}\|_\infty^3}{\kappa^3} > 0, \quad \text{when } \|\tilde{\delta}\|_\infty \text{ is small}. \quad (4.36)
$$
Hence (4.7) holds as well and Theorem 4.1 follows.

There now remains to prove (4.27). Let us denote by $g_\nu(e, \cdot)$ the Green function of the random walk with jump distribution $p_\nu(\cdot)$ and by $\phi_\nu(\cdot)$ the characteristic function of $p_\nu(\cdot)$. Then for $|e| = 1$ or $0$,

$$
P^\nu_e[H_0 < \infty] = \frac{g_\nu(e, 0)}{g_\nu(0, 0)} \quad \text{and} \quad \text{(4.37)}$$

$$
g_\nu(e, 0) = \int_T \frac{e^{-it\cdot e}}{1 - \phi_\nu(t)} \frac{dt}{(2\pi)^d}, \quad \text{with } t = (t_1, \ldots, t_d) \in T = (-\pi, \pi)^d. \quad \text{(4.38)}$$

Using the symmetry of $\phi_\nu$, we find:

$$
\frac{\partial}{\partial \nu} g_\nu(e, 0, 0) = \int_T \frac{\partial \phi_\nu}{\partial \nu} e^{-it\cdot e} \frac{dt}{(1 - \phi_\nu(t))^2 (2\pi)^d} = \int_T \frac{\partial \phi_\nu}{\partial \nu} \cos(t\cdot e) \frac{dt}{(1 - \phi_\nu(t))^2 (2\pi)^d} \quad \text{(4.39)}
$$

Note in particular that $\frac{\partial}{\partial \nu} g_\nu(0, 0)|_{\nu=1} = 0$, so that by (4.37)

$$
\frac{\partial}{\partial \nu}(P^\nu_{e,\nu}e[H_0 < \infty] - P^\nu_{e,\nu-1}[H_0 < \infty])|_{\nu=1} = -\frac{1}{g_{\nu=1}(0, 0)} \int_T \frac{(\cos t_{d-1} - \cos t_d)^2}{d(1 - \phi_\nu(1))^2} \frac{dt}{(2\pi)^d} < 0. \quad \text{(4.40)}
$$

On the other hand $P^\nu_{e,\nu-1}[H_0 < \infty] - P^\nu_{e,\nu-1}[H_0 < \infty] = 0$, by symmetry, and the claim (4.27) follows. □

**Remark 4.5.** – We know from Lawler [7], that for $\mu_0^\otimes \mathbb{Z}_d$-a.e. $\omega$, $P_0^\nu_{e,\nu} \omega$-a.s. $\frac{1}{\sqrt{n}}X_{[n]}$ converges in law to a Brownian motion with diagonal covariance matrix $A = \text{diag}(a_i)$, where

$$
a_i = 2 \int_{\Omega_0} \omega(0, e_i) dQ(\omega), \quad \text{for } 1 \leq i \leq d, \quad \text{(4.41)}
$$

and $Q$ is the unique invariant measure for the Markov chain of the environment viewed from the particle, which is absolutely continuous with respect to $\mu_0^\otimes \mathbb{Z}_d$. The measure $Q$ is known to be an ergodic invariant measure and from (1.40), we see that $\omega(0, \cdot)$ under $Q$ has same law as $\sigma_0^\nu$ under $Q^0$. As a by-product of the above example, cf. the choice (4.30), we see that one cannot in general replace the dynamic measure $Q$ with the static measure $\mu_0^\otimes \mathbb{Z}_d$ when calculating the limiting diffusion coefficient in (4.41).

**5. Perturbations of one-dimensional RWRE and velocity reversal**

We construct in this section another class of examples of multidimensional walks that satisfy the law of large numbers with a velocity which has an opposite direction to the expected local drift, or can vanish even if the latter does not vanish. The examples
in this section can be considered as perturbations of one-dimensional random walks in random environment, as opposed to the examples in Section 4 which were obtained as perturbation of the simple random walk in dimension $d$.

It is useful to first recall some known facts about one-dimensional random walks in random environment. Let $\mu$ denote a Borel probability measure on $(0,1)$, set $\Omega := (0,1)^Z$, and define the measure $\overline{P} = \mu^{\otimes Z}$ on the environment $\Omega$. For every $\tilde{\omega} \in \Omega$, the one-dimensional walk $\overline{X}_n$ under the law $\overline{P}_0 = \overline{P} \times P_{0,\tilde{\omega}}$ is defined as in $(0.3)$. Set $\rho_\varepsilon = (1 - \tilde{\omega}_2)/\tilde{\omega}_2$, define $d_0 = 2E_\overline{P}(\tilde{\omega}_0) - 1$ and $t_0 = E_\overline{P}(\log \rho_\varepsilon)$. The following facts are well known:

**Lemma 5.1.** 1) If $t_0 > 0$ then $\overline{P}_0$-a.s., $\lim \overline{X}_n = -\infty$. Further, if in addition there exists a constant $\kappa > 0$ such that $\mu[\tilde{\omega}_0 \in (\kappa, 1 - \kappa)] = 1$, then $E_0(\overline{X}_n) < 0$ for all $n$ large enough.

2) One may construct a law $\overline{\mu}$ with $d_0 > 0$, $\kappa > 0$, but $t_0 > 0$.

**Proof.** 1) That $\lim \overline{X}_n = -\infty$ is a consequence of [10]. Next, an application of [13, Proposition 2.6] shows that if $t_0 > 0$ and $\kappa > 0$ then, with $L_0 = \max\{X_n : n \geq 0\}$, it holds that for some constant $c_2 > 0$,

$$\overline{P}_0(L_0 > k) \leq \exp(-c_2 k). \tag{5.1}$$

In particular, $X_n \not\equiv 0$ is dominated by $L_0$. Since $t_0 > 0$ implies that $X_n \rightarrow -\infty$, $\overline{P}_0$-a.s., the above yields that $E_0(X_n) \rightarrow -\infty$, completing the proof of the first part of the lemma.

2) Take $\delta \in (0,1)$ small enough such that

$$\frac{1}{5} \log \frac{1-\delta}{\delta} - \frac{4}{5} \log 2 > 0,$$

and define $\overline{\mu}(\{\delta\}) = 1/5$ and $\overline{\mu}(\{2/3\}) = 4/5$. \square

Fix a $\tilde{\mu}$ as in part 2 of Lemma 5.1, and an $\varepsilon_0 > 0$ small enough such that, if $G_\varepsilon$ denotes a modified geometric random variable of parameter $\varepsilon$ independent of $\{\overline{X}_n\}$, then

$$A_0 := E_{0}(\overline{X}_{G_{\varepsilon}}) < 0 \tag{5.2}$$

(this is always possible due to part 1 of Lemma 5.1). For every $1 > \varepsilon \geq \varepsilon_0$ and $d_1 \geq 5$, set $d_2 = 1$, $q(e) = \varepsilon/2d_1$, $e \in \mathbb{Z}^{d_1}$, and $\mu \in \mathcal{P}_{q(e)}$ such that $\mu$ governs the law of the single site jump distribution conditioned on non-vanishing of the $\mathbb{Z}^{d_2}$-component. Let $X_n$ denote the random walk in random environment corresponding to the law $\overline{P} = \mu^{\otimes \mathbb{Z}}$, and let $v = v(\mu, d_1, \varepsilon)$ be the limiting velocity appearing in Theorem 1.4. Note that $v \cdot e = 0$ for every $e \in \mathbb{Z}^{d_1}$. Let $v_2 = v \cdot e_d$ denote the projection of $v$ into the direction corresponding to the $\mathbb{Z}^{d_1}$ subspace. We now claim the following:

**Theorem 5.2.** Fix $\tilde{\mu}$ and $\tilde{\varepsilon}_0$ as above. Then, there exists an integer $\tilde{d} = \tilde{d}(\tilde{\mu}, \tilde{\varepsilon}_0)$ such that for any $d_1 > \tilde{d}$, it holds that $v_2(\tilde{\mu}, d_1, \varepsilon_0) < 0$ while $\lim_{\varepsilon \rightarrow 1} v_2(\tilde{\mu}, d_1, \varepsilon)/(1 - \varepsilon) = d_0 > 0$.

By the continuity of $v_2(\tilde{\mu}, d_1, \varepsilon)$ in $\varepsilon$, which follows from similar considerations as in $(4.10)$, we see that for every $d_1 > \tilde{d}$ one may find an $\varepsilon > \varepsilon_0$ such that $v(\tilde{\mu}, d_1, \varepsilon) = 0$. 

Moreover, when \( d_1 > \tilde{d} \lor 13 \), Theorem 2.2 implies that the corresponding walk \( X_n \) exhibits a diffusive behavior.

It is interesting to comment on the nature of the phenomenon described in Theorem 5.2: for \( \varepsilon \) close to 1, between consecutive cut points of the \( \mathbb{Z}^{d_1} \) walk, \( X_n \) does not spend much time moving in the \( d \)th direction, and with high probability makes at most one step in that direction. This then averages out to give a positive displacement since \( d_0 > 0 \). On the other hand, when \( d_1 \) is large, most moves in the \( \mathbb{Z}^{d_1} \)-walk are cut points. If also \( \varepsilon \) is small enough, the walker effectively executes in the \( d \)-direction a one-dimensional random walk in random environment between cut points, for a geometric time of mean \( 1/\varepsilon \). That one-dimensional random walk in random environment is constructed such that while it does not have a negative speed (this is impossible since \( d_0 > 0 \)), it is transient to \(-\infty \) and hence leads to a negative displacement.

**Proof.** – Recall the cut times \( T' \). From (1.41) and similar considerations as in (4.9),

\[
\nu_2(\tilde{\mu}, d_1, \varepsilon_0) = \frac{E\tilde{Q}_\varepsilon[Z_{T'}^1, e_d]}{E^P[T']}
\]

Hence, the first part of the theorem follows as soon as we show that for \( \varepsilon = \varepsilon_0 \) and \( d_1 \) large enough it holds that

\[
E\tilde{Q}_\varepsilon[Z_{T'}^1, e_d] < 0. \tag{5.3}
\]

Define \( \mathcal{J} = \{ n : X_{k_n}^1 \neq X_{n-1}^1 \} \), and let \( \cdots < j_{-1} < j_0 < j_1 < \cdots \) denote the elements of \( \mathcal{J} \). Set \( V_n^1 = X_{j_n}^1 \), and note that under \( P = \delta_{000} \), \{ \( V_n^1 \) \} is a \( d_1 \)-dimensional simple random walk, independent of the i.i.d., geometric(\( \varepsilon_0 \)) random variables \( (j_{i+1} - j_i)_{i \in \mathbb{Z}^0}, j_1, -j_0 + 1 \). Recall the cut times \( T' \), note that \( T' \in \mathcal{J} \), and write \( J_i = j_i \), for the element of \( \mathcal{J} \) corresponding to \( T' \). Note that the \( c_i \) are precisely the cut times for the walk \{ \( V_n^1 \) \}.

Call a cut time \( T' \) **good** if \( X_n^1 = X_{T'}^1 \) for \( n \in [T', T' + 1 - 1] \), that is if \( J_{i+1} = j_{c_i+1} \). To prove (5.3), note first that

\[
E\tilde{Q}_\varepsilon[Z_{T'}^1, e_d] = E\tilde{Q}_\varepsilon[Z_{T'}^1, e_d1_{T' \text{ is good}}] + E\tilde{Q}_\varepsilon[Z_{T'}^1, e_d1_{T' \text{ is good}}] =: A + B.
\]

We claim that under the measure \( \hat{P}[\cdot \mid T^0 \text{ is good}] \), \( T' \) is geometric (\( \varepsilon_0 \)). Indeed, with \( D^V = \{ c_i \}_{i \in \mathbb{Z}} \) denoting the cut times of \{ \( V_n^1 \) \},

\[
P[T' = k, 0 \in D, T^0 \text{ is good}] = P[0 \in D^V, 1 \in D^V, j_0 = 0, j_1 = k]
\]

\[
= P[0 \in D^V, 1 \in D^V](1 - \varepsilon_0)^{k-1} \varepsilon_0^2,
\]

implying that

\[
\hat{P}[T' = k \mid T^0 \text{ is good}] = (1 - \varepsilon_0)^{k-1} \varepsilon_0.
\]

On the other hand, under the law \( \tilde{Q}_\varepsilon \), on the event \{ \( T^0 \text{ is good} \) \}, \( X_n^2 \) performs, for \( n \in [0, T' - 1] \), a one dimensional random walk in random environment, with environment generated by \( \tilde{\mu} \) (cf. (1.22)). Hence,
The estimate (5.4) and the Cauchy–Schwarz inequality imply then that
\[ P[T^0 \text{ is good}] \]

because (see [5], Remark 3, p. 248)
\[ P[T^0 \text{ is good}] = P[T^0 \text{ is good}. T_1 = k] \]

while, similarly,
\[ P[T^0 \text{ is good}] \]

Choosing as can be checked via characteristic functions, shows that, as a function of \( d \)

is uniformly bounded below for \( d \geq 5 \). Thus, \( A \to A_0 < 0 \). On the other hand,
a repeat of the proof of (1.12), using the fact that \( P[X_1^0 = 0] \) decreases with \( d \)
as can be checked via characteristic functions, shows that, as a function of \( d \geq 9 \),

is uniformly bounded. Hence, \( E \hat{\mathcal{Q}}[\{(T^1)^2\}] \) is uniformly bounded for \( d \geq 9 \).

The estimate (5.4) and the Cauchy–Schwarz inequality imply then that

\[ |B| \leq E \hat{\mathcal{Q}}[T^1 \mathbf{1}_{[T^0 \text{ is not good}]}] \to d_1 \to \infty \ 0. \]

Choosing \( d_1 \) large enough such that \( A + B < 0 \), the first part of the theorem follows.

The second part is actually easier: with the notations of (1.2),

\[ E \hat{\mathcal{Q}}[Z_{T^1} \cdot e_d] = \hat{\mathcal{Q}}[[\{n \in [1, T^1]: I_n = 0\} \mid = 1]d_0 \]

\[ + E \hat{\mathcal{Q}}[[\{n \in [1, T^1]: I_n = 0\} \mid \{n \leq 1\} + 1]Z_{T^1} \cdot e_d] =: d_0 C + D. \]

But, setting \( \tilde{j}_i = j_i \) for \( i \geq 1 \) and \( \tilde{j}_0 = 0 \),

\[ \hat{P}[[\{n \in [1, T^1]: I_n = 0\} \mid = 0] \]

\[ = \frac{P[\sum_{i=0}^{n-1} (\tilde{j}_{i+1} - \tilde{j}_i) = 0; \ j_0 = 0; \ 0 \in D^V]}{P(0 \in D)} = E \hat{\mathcal{P}}[\varepsilon^{c_1}], \]

while, similarly,

\[ \hat{P}[[\{n \in [1, T^1]: I_n = 0\} \mid > 1] \]

\[ \leq \hat{P}[\exists 0 < i < k \leq c_1 - 1: \ \tilde{j}_{i+1} - \tilde{j}_i = 1, \ \tilde{j}_{k+1} - \tilde{j}_k - 1 = 1] \]

\[ + \hat{P}[\exists 0 < i \leq c_1 - 1: \ \tilde{j}_{i+1} - \tilde{j}_i - 1 \geq 2] \]

\[ \leq (1 - \varepsilon^2)E \hat{\mathcal{P}}[(c_1)^2 + c_1]. \]
Note that the law $\hat{P}[c_1 \in \cdot]$ does not depend on $\varepsilon$. Since for $d_1 \geq 7$ it holds that $E\hat{P}(c_1^2) < \infty$, we conclude that $D/(1 - \varepsilon) \to \varepsilon \to 1^0$. Further, we also get

$$\lim_{\varepsilon \to 1} \frac{C}{1 - \varepsilon} = \lim_{\varepsilon \to 1} \frac{1 - E\hat{P}(c_1^{\varepsilon 1})}{1 - \varepsilon} = E\hat{P}[c_1].$$

Since also $\lim_{\varepsilon \to 1} P(c_1 \neq T^1) = 0$, one has that $E\hat{P}(T^1) \to \varepsilon \to 1^0 E\hat{P}(c_1)$, and the theorem follows. $\square$

**Remark 5.3.** – One easily adapts part 2 of Lemma 5.1 to construct a law $\bar{\mu}$ with $d_0 = 0$, $\kappa > 0$ and $\ell_0 > 0$ (take simply $\delta$ small enough with $\bar{\mu}(\{\delta\}) = 1/(4 - 2\delta)$ and $\bar{\mu}(\{2/3\}) = (3 - 2\delta)/(4 - 2\delta)$). A rerun of the proof of Theorem 5.2 then yields examples where the static expectation of the drift vanishes, but the limiting speed of the RWRE does not.

**REFERENCES**